

# Mandatory IFRS Adoption and Earnings Quality: Evidence from the UK

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Received: April 15, 2018 Accepted: September 20, 2018 Online Published: October 29, 2018

doi:10.5539/mas.v12n11p197

URL: <https://doi.org/10.5539/mas.v12n11p197>

## Abstract

This study examines the effect of mandatory IFRS adoption on the quality of accounting earnings for the firms listed in London Stock Exchange. After examining 9056 firm year observations for the period from 1994 to 2013 the results suggest that the mandatory adoption of IFRS leads to higher earnings quality. This study extends the current literature that examines the consequences of mandating IFRS adoption in the UK and shows that adopting high quality accounting standards leads to high quality accounting numbers.

**Keywords:** mandatory adoption, IFRS, earnings quality, accruals quality, UK

## 1. Introduction

Prior research suggests that mandatory adoption of IFRS leads to high-quality accounting numbers. The proponents of IFRS adoption argue that it improves the quality and quantity of financial disclosure (Ru& Baljit, 2018). This assertion is supported by empirical evidence that suggests IFRS adoption leads to high-quality earnings. For example, Houqe et al. (2012), and Barth et al. (2008) provide international evidence that IFRS adoption leads to higher quality accounting numbers. Consistent results from emerging market were provided by Ismail et al. (2013), where they find that IFRS adoption leads to higher earnings quality for Malaysian firms. In addition, other research documented an increase in the value relevance of accounting numbers following the adoption of IFRS. For instance, Devalle et al. (2010) find that the adoption of IFRS increases the value relevance of some adopting countries. Clarkson et al. (2011) document a decrease in the value relevance for the companies of Common Low Countries, while record an increase in the value relevance of accounting numbers for Code Low Countries, following the mandatory adoption of IFRS.

However, some researchers suggest that IFRS will increase managerial flexibility and discretion especially due to the lack of implementation guidance and poor enforcement (Ball et al., 2003; Leuz 2003; Ahmed, Neel, & Wang 2013). Moreover, (Barth et al., 2008) suggest that, IFRS is expected to restrict accounting alternatives which may reduce the ability of firms' accounting numbers to better reflect the firm's economic condition and performance. In addition, it is argued that the inherent flexibility in the principles-based IFRS standards may provide more opportunities for firms' managers to manage their numbers (Daske et al., 2008). Consequently, it is also likely that the IFRS adoption may result in more accounting manipulation and thus lower earnings quality to some extent. The financial reporting environment of The United Kingdom (UK) changed significantly after 2005, when the adoption of IFRS becomes mandatory. Such IFRS adoption may affect earnings quality primarily through additional disclosure and improved comparability, which, in turn, are expected to increase reporting transparency. The proponents of IFRS adoption argue that IFRS improve transparency by increasing the quantity and quality of financial disclosure.

This paper contributes to the literature that examines the economic consequences of mandatory IFRS adoption. Brüggemann et al. (2012) mention that, most of the literature on the consequences of mandatory IFRS adoption provide transitory evidence with low statistical power because of short history of IFRS adoption, also, they mention that most of these papers provide evidence from cross-country studies which make it difficult to

disentangle the effect of IFRS from other synchronous changes that may affect the financial reporting content. Following the recommendations of Brüggemann et al. (2012) this paper relied on data from one country for a long period after IFRS adoption to draw its results. Using a large sample of 9,049 firm year observations collected from 1385 distinct UK firms for the period from 1994 to 2013, the results suggest that mandatory IFRS adoption leads to increases in the earnings quality for the UK's firms. This result is robust after using different measure of earnings quality, and after using different tests models. Higher earnings quality numbers after the mandatory adoption of IFRS is in line with the view that higher quality accounting standards will lead to higher quality accounting numbers (Barth et al., 2008; Houque et al., 2012; Ismail et al., 2013).

In Section 2, we set out the theoretical framework for the study and develop the hypothesis. In Section 3, we describe the sample selection and the measurements of variables used in the study. In Section 4, we discuss the research results. In Section 5, we conclude the results of the study.

## 2. Prior Research and Hypothesis Development

To address why IFRS adoption could lead to higher earnings quality, we turn to the large and growing body of literature that provide imperial evidence of the consequences of IFRS adoption. The proponents of IFRS adoption argue that IFRS improve transparency by increasing the quantity and quality of financial disclosure. The current combination of IASs and IFRSs has generated more than 2000 disclosure requirements which preparers need to interpret and implement (Ernst & Young, 2006). Further, these requirements represent twice the number of standards that were required under UK GAAP and four times those that had been required in France, which leads to increases of up to 30 per cent in the length of post-IFRS adoption annual reports for a sample of EU firms (Ernst & Young, 2006). Moreover, Ball (2006) suggest that IFRS provides more accurate and timely financial statement information than any national standards, including the local standards of EU countries.

Consistent with the assertion that IFRS adoption improves the quality of financial disclosure; previous research finds that IFRS adoption has favorable capital market consequences. Barth et al., (2008); Devalle et al., (2010); Landsman et al. (2012); Tsalavoutas et al., (2012); Ismail et al., (2013); suggest an increase in the value relevance and information content of accounting numbers following the adoption of IFRS. In addition, Ballas et al., (2010); Barth et al., (2008); Doukakis, (2010); Houque et al., (2012); Ismail et al., (2013), Aksu & Espahbodi, (2012); and Daske & Gebhardt, (2006) find that IFRS leads to high quality accounting numbers and improved disclosure quality. Similarly, Horton et al., (2013); and Houque et al., (2014), provide evidence that IFRS adoption improve information environment through increasing analysts' forecast accuracy and reducing analysts' forecast dispersion. A reduction in information processing cost, crash risk, and cost of equity capital, following IFRS adoption are documented by (Armstrong et al., 2010; Daske et al., 2008; DeFond et al., 2011), (Mark L. DeFond et al., 2015), and (Daske et al., 2008; Kim, Shi, et al., 2014; Li, 2010; Palea, 2009), respectively. While increases in the informativeness of stock prices and the transparency and comparability of financial disclosure are suggested by (Beuselinck et al., 2010; Bissessur & Hodgson, 2012; Gillberto Loureiro & Taboada, 2012; Kim & Shi, 2012a); and (Brochet et al., 2013).

However, the opponents of IFRS argue that IFRS has increased managerial flexibility and discretion especially due to the lack of implementation guidance and poor enforcement (Ball et al. 2003; Leuz 2003; Ahmed, Neel, & Wang 2013). In, addition, limiting accounting alternatives may reduce the ability of firms' accounting numbers to reflect the firm's economic condition and performance (Barth et al., 2008). In addition, it is argued that the inherent flexibility in the principles-based IFRS standards may provide more opportunities for firms' managers to manage their numbers (Daske et al., 2008). Consequently, it is also likely that the IFRS adoption may result in more accounting manipulation and thus lower earnings quality to some extent.

Capkunet al. (2012) report an increase in earnings management from pre-2005 to post-2005 for early and late voluntary adopters in countries that allowed early IFRS adoption, and for mandatory adopters in countries that did not allow early IFRS adoption. Similarly, Ahmed et al. (2013) provide evidence suggesting that following the mandatory adoption of IFRS, income smoothing and accrual aggressiveness increase while timeliness of loss recognition decreases. Christensen et al. (2015) examine the effect of IFRS adoption on income smoothing for German firms and find that income smoothing decreased significantly under voluntary but not under mandatory IFRS adoption. They suggest that mandatory adopters might perceive fewer benefits from principles based set of accounting standards (i.e. IFRS) and thus avoid the costs of transferring to IFRS. Paananen (2008) and Paananen & Lin (2009) both find a decrease in financial reporting quality, an increase in earnings management, and a reduction in timeliness of loss recognition, following IFRS adoption in Germany. Jeanjean & Stolowy (2008) find no decline in the pervasiveness of earnings management in Austria and the United Kingdom and find an

increase in France. Both Ahmed et al. (2013) and Chen et al. (2010) find evidence of income smoothing and a reduction in timeliness of loss recognition following mandatory IFRS. However, contrary to Chen et al. 2010, Ahmed et al. (2013) find a significant increase in aggressive reporting of some accruals and no reduction in the management of earnings towards a target. Prior studies therefore provide mix evidence about the net effect of IFRS adoption in earnings quality.

These mixed evidences could be explained by that some IFRS voluntary adopters may choose to adopt IFRS name without making any serious reporting changes (Daske et al., 2013). The fact that the results of the literature that examined the consequences of mandatory IFRS were driven from short history of mandatory IFRS adoption, reflecting a combination of idiosyncratic, temporary effects of first-time adoption and low statistical power due to relatively short analysis periods, two to three years after the mandatory adoption (Brüggenmann et al., 2013), could explain also the inconsistent results of IFRS literature. As the net effect of IFRS adoption is ambiguous, we will re-examine the effect of IFRS on earnings quality by testing the following null hypothesis:

*H0: The mandatory adoption of IFRS did not significantly affect the earnings quality of the UK firms.*

### 3. Sample and Methodology

In this section, we describe our sample and the methodology used to test the hypothesis developed above. Section 3.1 briefly describes our data sources and the procedures followed in sample selection process, and Section 3.2 provides an overview of our design, along with a detailed description of our variables. We use a fixed effect design to test our predictions on the effect of mandatory IFRS adoption on earnings quality.

#### 3.1 Data

The initial research sample consists of all the firms listed in London Stock Exchange that have available data in DataStream, Worldscope, and IBES international databases for the period between 1<sup>st</sup> January 1994 and 31<sup>st</sup> December 2013 (The three databases are combined in one software called DataStream published by Thomson Reuter). The sample period was chosen to cover the periods before the adoption and after the adoption, thus the effect of IFRS adoption on the earnings quality can be examined. We took long time period for two reasons. First, choosing a long period before and after IFRS adoption could result in a better measure of the effect of IFRS adoption on earnings quality. Since taking a long period before the adoption provides a better measure of the average earnings quality levels before the mandatory adoption of IFRS. Second, Wang & Yu (2015) suggest that one of the advantages of taking a large sample for a longer time period is that the results and conclusions, drawn from this sample, are more representative. The investigation is undertaken in one country, the UK, to hold constant certain institutional factors such as stock listing requirements, accounting disclosure requirements, market microstructures and regulatory environments that may confound the results, thereby strengthening the reliability of our findings, as suggested by Ruland et al. (2007), and Paananen & Lin (2009). In addition, according to Schipper (2005) choosing one country instead of multi countries helps in minimizing the heterogeneity and the cross-countries differences that may have an effect on the dependent variable.

Only firms listed in the L.S.E, active and dead, with available data for the period from 1990 to 2013 were included in the sample. Following the prior research (Hutton et al., 2009; Kim & Shi, 2012) the firms in financial, banking, and insurance industries, with SIC code 6000-6999 were excluded from the sample, because these industries have special regulations and financial accounting standards and the inclusion of these industries in the sample may distort the research results. Any firms with unavailable data to calculate the explanatory variables, and the dependent variable were excluded from the sample also. After applying the previous procedures, the final sample consists of 9049 firm-year observations collected from 1385 distinct UK firms for the period from 1994 to 2013.

#### 3.2 Methodology and Variable Measurement

We estimate fixed effect regression model and perform robustness analyses to test the hypothesis as follows:

To test H1 we specifying the following regression model:

$$MJ\_Model_{i,t} = \alpha_0 + \beta_1 IFRS_{i,t} + \beta_2 LEV_{i,t} + \beta_3 M\backslash B_{i,t} + \beta_4 SIZE_{i,t} + \beta_5 ROA_{i,t} + Industryfixedeffect + \varepsilon_{i,t} \quad (1)$$

In our robustness test we used different measure of earnings quality and estimate different regression models to examine the effect of IFRS adoption on earnings quality.

#### 3.3 Earnings Quality Measures

Given the fact that, the accruals quality models have been shown to be the most popular model in measuring

earnings quality, where (Dechow et al., 2010) reviewed more than 300 papers on earnings management determinants and consequences, and claims that the Jones (1991) Model, and the Modified Jones (1995) Model, are the top two in the list of the most commonly used measures of earnings quality. The measurement of earnings quality related to accruals quality is used in this study. Using accruals quality as a measure of earnings quality is consistent with previous research in this area, including Rajgopal and Venkatachalam (2011), Mouselli, Jaafar, & Hussainey (2012), Ismail et al. (2013), and Doukakis (2014), Watrin & Ullmann (2012), Hutton et al. (2009) and Kothari et al. (2005).

Following Kothari et al. (2005), the equation for nondiscretionary accruals for the Modified Jones Model (1995) is expressed as follows:

$$NDACC_{i,t} = \alpha_{i,t} 1/LTA + \alpha_1 (\Delta REV_{i,t} - \Delta REC_{i,t}/LTA_{i,t}) + \alpha_2 PPE_{i,t}/LTA_{i,t} \quad (2)$$

Where

$NDACC_{i,t}$  = non-discretionary accruals for firm  $i$  in year  $t$ .

$LTA_{i,t}$  = lagged total assets for firm  $i$  in year  $t$ .

$\Delta REV_{i,t}$  = change in revenues for firm  $i$  in year  $t$ .

$\Delta REC_{i,t}$  = change in account receivable for firm  $i$  in year  $t$ .

$PPE_{i,t}$  = property, plant and equipment for firm  $i$  in year  $t$ .

To calculate the nondiscretionary accruals using the Modified Jones Model (1995), it is necessary to estimate the coefficients  $\alpha_{i,t}$ ,  $\alpha_1$ , and  $\alpha_2$ , for the above model. The ordinary least squares regression was used to estimate the coefficients for each industry in each year to partially control for industry level changes in economic conditions, that affect total accruals and allows the coefficients to vary across time (Doukakis, 2014)

Consistent with the approach used by Athanasakou, Strong, & Walker (2009), industries with less than six observations in each year were removed from the sample, because of the lack of quorum in calculating the coefficient. The industry classification was based on the two digits SIC code classification.

In order to obtain the coefficients for the model in e.q (1), we estimate the following cross-sectional regression model using the firms in each two digit SIC code for each year between 1994 and 2013:

$$\frac{TA_{i,t}}{LTA_{i,t}} = \alpha \left( \frac{1}{LTA_{i,t}} \right) + \alpha_1 \left( \frac{\Delta REV_{i,t}}{LTA_{i,t}} \right) + \alpha_2 \left( \frac{PPE_{i,t}}{LTA_{i,t}} \right) + \varepsilon_{i,t} \quad (3)$$

Where:

$TA_{i,t}$  = total accrual for firm  $i$  in year  $t$ .

$LTA_{i,t}$  = lagged total asset for firm  $i$  in year  $t$ .

$\Delta REV_{i,t}$  = change in revenues for firm  $i$  in year  $t$ .

$PPE_{i,t}$  = property, plant and equipment for firm  $i$  in year  $t$ .

The coefficient from this regression model is used to calculate the nondiscretionary accruals (NDACC) based on the Modified Jones (1995) Model and the Jones (1991) model. Finally, the discretionary (abnormal) accruals calculating the difference between total accruals and the fitted normal accruals as follows:

$$AAC_{i,t} = (TA_{i,t}/LTA) - NDACC_{i,t} \quad (4)$$

All the variables are divided by lagged total assets to reduce heteroscedasticity (Jones, 1991).

Where:

$AAC_{i,t}$  = abnormal accruals for firm  $i$  in year  $t$ .

$TA_{i,t}$  = total accrual for firm  $i$  in year  $t$ .

$LTA_{i,t}$  = lagged total asset for firm  $i$  in year  $t$ .

$NDACC_{i,t}$  = non-discretionary accruals for firm  $i$  in year  $t$ .

The earnings quality for each firm is estimated through the absolute value of the abnormal accruals ( $|AAC_{i,t}|$ ). The large quantity of absolute value of abnormal accruals indicates low earnings quality and vice versa.

### 3.4 IFRS Definition

The information about the accounting standards followed in preparing the firm's financial statements was obtained

from DataStream database. The DataStream code (WC07536) provides information about the accounting standards followed in preparing the financial statements for a specific firm. Table 1 provides a detailed description of the Worldscope code (WC07536) classification of the accounting standards followed by each firm. DataStream identifies 23 different accounting standards that are used by firms to prepare the financial statements. This identification ranges from local accounting standards (07536 = 1), International Accounting Standards (IAS) pronounced by International Accounting Standards Committee IASC (07536 = 2), U.S. standards (07536 = 3), accounting standards that adopt local standards with other gridlines (07536 = 08, 10, 17), or other hybrid type accounting standards that adopt local standards along with international accounting standards (07536 = 18,19). The information about Worldscope accounting standards classifications is retrieved from Thomson Reuters (2012) website. This research follows Kim & Shi (2012a), by identifying the firm as an IFRS adopter, if it adopts a full set of IFRS or IAS (07536 = 02 or 23), and marked as a non-adopter if it adopts any other accounting standards. In particular, if the firm adopts IAS or IFRS with another set of accounting standards, then this firm is considered as a non-adopter.

### 3.5 Control Variables

We include in the regression model several control variables that previous literature finds or suggests a relationship between these variables and earnings quality. Firm size, as measured by firm's total asset, is expected to influence earnings quality. McNinnis & Collins (2011) expect that firm size could have a negative effect on earnings quality.

#### 3.5.1 Size

Watts & Zimmerman (1978) suggest that large firms face greater political costs relative to small firms since larger firms are expected to be followed by more financial analysts and investors due to their larger market value and, hence, greater influence on the market. If this is true, then larger companies would have a greater incentive to reduce earnings. In addition, Lobo & Zhou (2006) note that large firms may have more opportunities to overstate earnings because of the complexity of their operations and the difficulty for users to detect overstatement. Either way, size is a variable that could potentially bias the coefficients of the variables of interest in the study. Hence, we include a size variable to control for potential earnings overstatement.

Table 1. Worldscope description of accounting followed (Field 07536)

Worldscope fields 07536	Worldscope description
1	Local standards
2	International standards
3	U.S. standards (GAAP)
4	Commonwealth countries standards
5	EU standards
6	International standards and some EU guidelines
7	Specific standards set by the group
8	Local standards with EU and IASC guidelines
9	Not disclosed
10	Local standards with some EU guidelines
11	Local standards – inconsistency problems
12	International standards – inconsistency problems
13	US standards – inconsistency problems
14	Commonwealth standards – inconsistency problems
15	EEC standards – inconsistency problems
16	International standards and some EU guidelines – inconsistency problems
17	Local standards with some OECD guidelines
18	Local standards with some IASC guidelines

19	Local standards with OECD and IASC guidelines
20	US GAAP reclassified from local standards
21	Local standards with a certain reclassification for foreign companies
22	Other
23	IFRS

### 3.5.2 Growth Opportunity

Following the outcomes of An & Zhang (2013), Bae et al. (2013), Hasan et al. (2014), He et al. (2013), Gul et al. (2010), and Hutton et al. (2009) the growth opportunity is calculated as the ratio of market value of equity to book value of equity. Summers & Sweeney (1998) note that unethical managers may be induced to misstate financial statements when growth slows or reverses, in order to maintain the appearance of consistent growth. The market to book ratio represents market expectations of future profitability growth. Trying to meet such expectations could be a major motive for managers and discretionary accruals is a useful tool to attain that objective. McNnis & Collins (2011) suggest that firms with higher market-to-book ratios (MTB) have greater incentives to meet earnings targets, so we expect MTB to have negative effect on earnings quality.

### 3.5.3 ROA

Firm's performance and profitability, as measured by the ratio of net income to total assets (Ben-Nasr & Cosset, 2014; and Gul, Srinidhi, et al. 2011) is believed to have an effect on earnings quality. We include ROA in the regression to control for differences in performance between firms across the sample because prior research documents a relation between abnormal accruals and performance (e.g., Dechow et al., 1995; McNnis & Collins, 2011). Additionally, Kothari et al. (2005) offer evidence that controlling for ROA reduces the probability of Type I error in earnings quality studies where performance differences are not part of the hypotheses being tested.

### 3.5.4 Leverage

Leverage is measured as the ratio of total liabilities to total assets and is included to control for the possible effects (positive or negative) of firm risk. Velury & Jenkins (2006) and Eng & Mak (2003) find that managerial ownership and debt (firm size) are negatively (positively) related to disclosure quality. DeFond & Jiambalvo (1994) presents evidence that managers of highly leveraged firms have incentives to make income increasing discretionary accruals to avoid debt covenant violation. Based on the previous discussion we expect leverage to have positive effect on abnormal accruals.

## 4. Discussion of Results

### 4.1 Descriptive Statistics and Correlation Test

Table (2) presents descriptive statistics for the variables used in our analysis. The mean and median for the discretionary (abnormal) accruals, using Modified Johns model, are 6.36% and 3.95% of total assets, respectively. These values are comparable to that in earnings quality literature. Iatridis (2012) in his UK study find that the average value of discretionary accruals using Jones (1991) model is about 6%. The statistics for other variables is as follow; growth opportunity (as measured by market value of equity to book value of equity) range from -7.3 to 27.85 with a mean value 3.043. Leverage, size, and ROA record a mean (median) value of, 0.201 (0.172), 11.790 (11.566), and 0.013 (0.043), respectively.

Table 2 present the descriptive statistics for the variable of interest. The statistics are computed from 9049 firm year in the sample period 1994-2013

	Num. of Obs	MEAN	STD.DEV.	MIN	MAX	median
ABS_JOHNS	9049	0.062	0.0668	0.0002	0.31	0.039
ABS_MJOHNS	9049	0.064	0.0725	0.001	0.349	0.039
GROWTH_OPPURTUNITY	9049	3.043	4.527	-7.3	27.85	1.93
LEVERAGE	9049	0.201	0.1841	0.0001	3.714	0.172
SIZE	9049	11.79	2.2204	0	19.414	11.566
ROA	9049	0.013	0.2401	-1.308	0.943	0.043

IFR	9049	0.577	0.494	0	1
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Table (3) Presents Pearson correlation matrix for the dependent (earnings quality), independent (IFRS adoption) and all the control variables used in the regression analysis. The Pearson correlation coefficient is a measure of the strength of the linear relationship between two variables. The highest correlation coefficient is between growth opportunity and leverage, and the lowest correction is between ROA and IFRS. The correlation between IFRS adoption and abnormal accruals is negative suggesting that the IFRS adoption could lead to higher earnings quality. Because the correlation coefficients for all the variables in the correlation analysis matrix are below 80%, we can suggest that there is no multicollinearity problem in the data (Gujarati & Porter, 2009; Hair et al., 2010).

Table 3. Reports the correlation coefficient between key variables; the correlations are computed from 9049 firm year in the sample period 1994-2013.

	ABS_MJONS	GROWTH_OPPORTUNITY	LEVERAGE	SIZE	ROA	IFRS
ABS_MJONS	1					
GROWTH_OPPORTUNITY	0.0507	1				
LEVERAGE	0.062	-0.716	1			
SIZE	-0.1404	-0.0704	0.1357	1		
ROA	-0.2594	-0.0875	-0.1049	0.3377	1	
IFRS	-0.038	-0.0977	0.0237	0.0819	-0.0064	1

#### 4.2 Main Results

Table 4 (Modified Johns model) Fixed effect model with adjusted standard error for hetroscedasticity.

Table 4. Multivariate regression analysis of IFRS adoption and earnings quality. The sample consists of 9049 firm year observations from 1385 UK listed firms for the period between 1994 and 2013. This regression results based on panel data industry fixed effect model. The dependent variable is earning quality calculated by *Modified Johns model*. The main independent variable is the mandatory adoption of IFRS; the first column presents the dependent variables. The second column presents the estimated coefficients change in the dependent variable as a result of one unit change in the independent variable. The third, fourth and fifth columns represent the standard errors, t\_test and p\_value results respectively. Here \*, \*\*, \*\*\* present 10, 5, 1 % levels of significant respectively for two tailed test

	coef.	std.err	t	p
GROWTH_OPPORTUNITY	0.0002	0.0002	0.89	0.372
LEVERAGE	0.0051	0.0076	0.67	0.5
SIZE	-0.0048	0.0018	-2.68	0.007***
ROA	-0.6382	0.0090	-7.07	0***
IFRS	-0.0072	0.0021	-3.47	0.001***
Const	0.1222	0.021	5.82	0***

The significant negative effect of IFRS variable on abnormal accruals is consistent with the assertion that IFRS adoption improves the quality of financial disclosure. Previous research finds that IFRS adoption has favorable capital market consequences. (Barth et al., 2008; Devalle et al., 2010; Ismail et al., 2013; Tsalavoutas et al., 2012; Landsman et al., 2012) suggest an increases in the value relevance and information content of accounting numbers following the adoption of IFRS. In addition, (Ballas et al., 2010; Barth et al., 2008; Doukakis, 2010; Houque et al., 2012; Ismail et al., 2013, Aksu & Espahbodi, 2012; Daske & Gebhardt, 2006) find that IFRS leads to high quality accounting numbers and improved disclosure quality. Similarly, (Horton et al., 2013; Houque et al., 2014), provide evidence that IFRS adoption improve information environment through increasing analysts'

forecast accuracy and reducing analysts' forecast dispersion.

Prior research suggests that mandatory adoption of IFRS leads to high-quality accounting numbers. For example, Houqe et al. (2012), and Barth et al. (2008) provide international evidence that IFRS adoption leads to higher quality accounting numbers. Consistent results from emerging markets were provided by Ismail et al. (2013). In addition, other research documented an increase in the value relevance of accounting numbers following the adoption of IFRS. For example, Devalle et al. (2010) and Clarkson et al. (2011) find that the adoption of IFRS increases the value relevance for some adopting countries. Based on previous discussion one can expect higher levels of earnings quality following the mandatory adoption of IFRS. There is a positive effect of leverage on discretionary accruals, this mean that high leverage leads to lower earnings quality. This result is consistent with the results of Velury & Jenkins (2006) and Eng & Mak (2003) who find that debt is negatively related to disclosure quality. DeFond & Jiambalvo (1994) presents evidence that managers of highly leveraged firms have incentives to make income increasing discretionary accruals to avoid debt covenant violation.

Consistent with the expectation, we find a positive association between MTB and discretionary accruals. This result is in line with McInnis and Collins (2011) who suggest that firms with higher market-to-book ratios (MTB) have greater incentives to meet earnings targets. Firms' size shows negative effect on abnormal accruals. This result suggests that big firms tend to have higher quality earnings numbers. The results is consistent with the results of Christensen et al. (2015), and Fang, Huang, & Karpoff (2016). The significant negative effect of ROA on discretionary accruals suggests that, more profitable firms have higher quality earnings numbers.

#### 4.2 Robustness Tests

Our results are robust after using different models and after using different measures of earnings quality. We repeat the regression by using pooled OLS regression, and fixed effect model with heteroscedasticity adjusted standard error and the results are robust. In addition, we repeat the regression by using earnings quality as measured by John's model, and using OLS, fixed effect model, and fixed effect model with heteroscedasticity adjusted standard error, and the results were consistent with our main regression results. All the tables for robustness test can be found in the appendix.

### 5. Summary and Conclusion

This study investigates the impact of the mandatory adoption of IFRS on the earnings quality of the UK firms. Using a sample of 9049 firm year observations collected from 1385 distinct firms, we run regression analyses to compare those firms' earnings quality between the pre- and post-IFRS periods. The regression results suggest that earnings quality, as measured by discretionary accruals did improved following the mandatory adoption of IFRS. This research documents clear and novel evidence on whether the mandatory adoption of IFRS improved the earnings quality of the UK firms. The study findings may provide implications to the users of accounting information and the standard setters. It is clear that the IFRS adoption has improved firm's accruals quality, so it could benefit accounting information users in this respect. The findings may also help the standard setters in the UK in considering whether or not to adopt IFRS for UK's private firms in the future. Hence, the results can serve as a piece of timely evidence for evaluating the IFRS adoption in the UK.

In addition, this paper contributes to the literature that examines the economic consequences of mandatory IFRS adoption. Most of the literature on the consequences of mandatory IFRS adoption provide transitory evidence with low statistical power because of short history of IFRS adoption Brüggemann et al. (2012). Also, most of these papers provide evidence based on multi-countries data which make it difficult to distinguish the effect of IFRS from other concurrent changes that may affect the financial reporting content. By using large sample for long time period, from one county, we followed the future research recommendations of Brüggemann et al. (2012).

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## Appendix A

### Robustness tests

#### OLS Regression

	coef.	std.err	t	p
GROWTH_OPPORTUNITY	0.00042	0.00016	2.57	0.01
LEVERAGE	0.01942	0.00409	4.75	0
SIZE	-0.00211	0.000359	-5.9	0
ROA	-0.06953	0.003303	-21.05	0
IFRS	-0.0048	0.001498	-3.21	0.001
Const	0.08525	0.00431	19.78	0

#### VIF RESULTS

Variable	VIF	1/VIF
Size	1.18	0.849529
ROA	1.17	0.855832
LEVERAGEZE~N	1.05	0.950574
Growth opportunity	1.02	0.976008
IFRS	1.02	0.982989
Mean VIF	1.09	

#### (Modified Johns Model) Fixed Effect Model

	coef.	std.err	t	P
GROWTH_OPPORTUNITY	0.00022	0.000195	1.14	0.252
LEVERAGE	0.005147	0.006573	0.78	0.434
SIZE	-0.0048	0.001399	-3.46	0.001
ROA	-0.6382	0.004378	-14.58	0
IFRS	-0.00717	0.001895	-3.79	0
Const	0.1222	0.01628	7.51	0

(Johns Model) Fixed Effect Model with Adjusted Standard Error for Heteroscedasticity

fe with robust std. err				
	coef.	std.err	t	p
GROWTH_OPPORTUNITY	0.00022	0.00023	0.99	0.323
LEVERAGE	0.00554	0.00696	0.8	0.426
SIZE	-0.00455	0.00169	-2.69	0.007
ROA	-0.05838	0.00804	-7.26	0
IFRS	-0.00678	0.00192	-3.52	0
Const	0.1171	0.0196	5.96	0

(Johns Model) Fixed Effect Model

	coef.	std.err	t	p
GROWTH_OPPORTUNITY	0.00022	0.000195	1.14	0.252
LEVERAGE	0.00514	0.006573	0.74	0.434
SIZE	-0.0048	0.001399	-3.46	0.001
ROA	-0.0638	0.004378	-14.58	0
IFRS	-0.00717	0.001896	-3.79	0
const	0.12225	0.01628	7.51	0

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