On the Existence Conditions for Balanced Fractional 2^m Factorial Designs of Resolution R*({1}| Ω_ℓ) with $N < \nu_\ell(m)$

Yoshifumi Hyodo^{1,2}, Masahide Kuwada², Hiromu Yumiba²

¹ Graduate School of Informatics, Okayama University of Science, Okayama, Japan

² International Institute for Natural Sciences, Kurashiki, Japan

Correspondence: Hiromu Yumiba, International Institute for Natural Sciences, Kurashiki, 710-0821, Japan. E-mail: yumiba@edu.kake.ac.jp

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Abstract

We consider a fractional 2^m factorial design derived from a simple array (SA) such that the $(\ell + 1)$ -factor and higher-order interactions are assumed to be negligible, where $2\ell \le m$. Under these situations, if at least the main effect is estimable, then a design is said to be of resolution $R^*(\{1\}|\Omega_\ell)$. In this paper, we give a necessary and sufficient condition for an SA to be a balanced fractional 2^m factorial design of resolution $R^*(\{1\}|\Omega_\ell)$ for $\ell = 2, 3$, where the number of assemblies is less than the number of non-negligible factorial effects. Such a design is concretely characterized by the suffixes of the indices of an SA.

Keywords: association algebra, balanced fractional factorial design, estimable parametric function, factorial effect, resolution, simple array

1. Introduction

As a generalization of an orthogonal array, the concept of a balanced array (BA) was first introduced by Chakravarti (1956) as a partially BA. However it is a generalization of the BIB design rather than of the PBIB design. Thus Srivastava and Chopra (1971) called it by BA. A BA of strength t, size N, m constraints, two symbols and index set $\{\mu_i^{(t)} | 0 \le i \le t\}$ is briefly written by BA(N, m, 2, t; $\{\mu_i^{(t)}\}$). In particular, a BA of strength t = m is called a simple array (SA) (see Shirakura, 1977), and it is written by SA(m; $\{\lambda_x\}$) for brevity, where $\lambda_x = \mu_x^{(m)}$. When t < m, a BA of strength t does not always exist for given indices $\mu_i^{(t)}$. On the other hand, an SA always exists for any λ_x and any *m*. The existence conditions for a BA of strength t were given by Srivastava (1972) for m = t + 1, t + 2, and Shirakura (1977) for m = t + 3. If the variance-covariance matrix of the estimators of the factorial effects to be of interest is invariant under any permutation on the factors, then a design is said to be balanced. Under certain conditions, a BA of strength 2ℓ turns out to be a balanced fractional 2^m factorial (2^m -BFF) design of resolution $2\ell + 1$ (see for $\ell = 2$, Srivastava, 1970, and for general ℓ , Yamamoto *et al.*, 1975), where $2\ell \leq m$. The characteristic roots of the information matrix of a 2^m -BFF design of resolution V, i.e., $\ell = 2$, were obtained by Srivastava and Chopra (1971). By using the triangular multidimensional partially balanced (TMDPB) association scheme and its algebra, their results were generalized by Yamamoto et al. (1976) and Hyodo (1992) for a resolution $2\ell + 1$ design, where $2\ell \le m$ and $m < 2\ell \le 2m$, respectively. The concept of the MDPB association scheme, which is a generalization of an ordinary association scheme (e.g., Bailey, 2004), was introduced by Bose and Srivastava (1964). The existence conditions for a BA of strength 2ℓ to be a 2^m -BFF design of resolution 2ℓ for general ℓ were obtained by Shirakura (1975,1980). Some algebraic properties of the information matrix of a fractional 2^m factorial (2^m -FF) design derived from an SA were investigated by Hyodo and Yamamoto (1988) and Hyodo (1989). As the extension of the concept of resolution, Yamamoto and Hyodo (1984) discussed the extended concept of resolution for 2^m fractions.

Definition 1.1. Under the assumption that the $(\ell + 1)$ -factor and higher-order interactions are negligible, if the p_1 -factor, the p_2 -factor, \cdots , and the p_r -factor interactions are estimable, where $0 \le p_1 < p_2 < \cdots < p_r \le \ell$, and furthermore if the

remaining interactions are not estimable (including the general mean and the main effect), then a design is said to be of resolution $R(\{p_1, p_2, \dots, p_r\}|\Omega_\ell)$, where $\Omega_\ell = \{0, 1, \dots, \ell\}$. In particular, when $p_i = i - 1$ ($1 \le i \le r = \ell + 1$), it is of resolution $2\ell + 1$, and when $p_i = i (1 \le i \le r = \ell - 1)$ (or $p_i = i - 1 (1 \le i \le r = \ell)$), it is of resolution 2ℓ .

By relaxing the conditions of Definition 1.1, we give the following definition of resolution:

Definition 1.2. Under the same assumptions as Definition 1.1, if at least the p_1 -factor, the p_2 -factor,..., and the p_r -factor interactions are estimable, where $0 \le p_1 < p_2 < \cdots < p_r \le \ell$, then a design is said to be of resolution $\mathbb{R}^*(\{p_1, p_2, \cdots, p_r\}|\Omega_\ell)$.

Note that the set of resolution $R(\{p_1, p_2, \dots, p_r\}|\Omega_\ell)$ designs is a subset of resolution $R^*(\{p_1, p_2, \dots, p_r\}|\Omega_\ell)$ designs. For example, a resolution $R^*(\{1\}|\Omega_3)$ design is of resolution $R(\omega|\Omega_3)$, where $\omega = \{1\}$, $\{0, 1\}$, $\{1, 2\}$, $\{1, 3\}$, $\{0, 1, 2\}$, $\{0, 1, 3\}$, $\{1, 2, 3\}$ or $\{0, 1, 2, 3\}$. Here when a design is derived from an SA, where the number of assemblies (or treatment combinations) is less than the number of non-negligible factorial effects, there does not exist a resolution $R(\{1, 2\}|\Omega_3)$, $R(\{1, 3\}|\Omega_3)$, $R(\{0, 1, 2, 3\}|\Omega_3)$ or $R(\{0, 1, 2, 3\}|\Omega_3)$ or $R(\{0, 1, 2, 3\}|\Omega_3)$ design (see Table 4.1 in latter).

In a practical experiment, the most interesting factorial effect may be the main effect, next may be the two-factor interaction, and so on. Using the algebraic structure of the TMDPB association scheme and the matrix equations, Kuwada *et al.* (2003) obtained a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_3)$. However their results are very complex. A necessary and sufficient condition for an SA to be a 2^m -BFF design of resolution $2\ell + 1$ for general ℓ has been obtained by Hyodo *et al.* (2015), where $2\ell \leq m$.

In this paper, we consider a 2^m -BFF design derived from an SA such that the number of assemblies is less than the number of factorial effects up to the ℓ -factor interaction, where $\ell = 2, 3$. Under these situations, using the suffixes of the indices λ_x of an SA, we give a necessary and sufficient condition for an SA to be a 2^m -BFF design of resolution R^{*}({1}|\Omega_2), i.e., of resolution IV, and also we rewrite a necessary and sufficient condition for an SA to be a 2^m -BFF design of resolution R^{*}({1}|\Omega_2), i.e., R^{*}({1}|\Omega_3).

2. Preliminaries

We consider a 2^m -FF design T with m factors and N assemblies, where $2\ell \le m$, and the $(\ell + 1)$ -factor and higher-order interactions are assumed to be negligible. Then the linear model is given by $\mathbf{y}(T) = E_T \boldsymbol{\Theta} + \boldsymbol{e}_T$, where $\mathbf{y}(T)$ is an $N \times 1$ observation vector, E_T is the $N \times v_\ell(m)$ design matrix, $\boldsymbol{\Theta}' = (\boldsymbol{\theta}'_0; \boldsymbol{\theta}'_1; \cdots; \boldsymbol{\theta}'_\ell)$, and \boldsymbol{e}_T is an $N \times 1$ error vector with mean $\boldsymbol{\theta}_N$ and variance-covariance matrix $\sigma^2 I_N$. Here $\boldsymbol{\theta}_0, \boldsymbol{\theta}_1, \cdots$, and $\boldsymbol{\theta}_\ell$ are the general mean, the vector of the main effect, \cdots , and the vector of the ℓ -factor interaction, respectively, $v_\ell(m) = \binom{m}{0} + \binom{m}{1} + \cdots + \binom{m}{\ell}$, and I_p is the identity matrix of order p. The normal equations for estimating $\boldsymbol{\Theta}$ are given by $M_T \hat{\boldsymbol{\Theta}} = E'_T \mathbf{y}(T)$, where $M_T = E'_T E_T$ is the information matrix of order $v_\ell(m)$. If M_T is non-singular, then T is of resolution $2\ell + 1$.

Let $A_{\alpha}^{(u,v)} (= A_{\alpha}^{(v,u)'})$ $(0 \le \alpha \le u \le v \le \ell)$ be the local association matrices of size $\binom{m}{u} \times \binom{m}{v}$ of the TMDPB association scheme, and further let $A_{\beta}^{\#(u,v)} (= A_{\beta}^{\#(v,u)'})$ $(0 \le \beta \le u \le v \le \ell)$ be the matrices of size $\binom{m}{u} \times \binom{m}{v}$ (see Yamamoto *et al.*, 1976), where the relation between $A_{\alpha}^{(u,v)}$ and $A_{\beta}^{\#(u,v)}$ is given by

$$A_{\alpha}^{(u,v)} = \sum_{\beta=0}^{u} z_{\beta\alpha}^{(u,v)} A_{\beta}^{\#(u,v)} \quad \text{for } 0 \le \alpha \le u \le v$$

and

$$A_{\beta}^{\#(u,v)} = \sum_{\alpha=0}^{u} z_{(u,v)}^{\beta\alpha} A_{\alpha}^{(u,v)} \quad \text{for } 0 \le \beta \le u \le v.$$

Here

$$z_{\beta\alpha}^{(u,v)} = \sum_{b=0}^{\alpha} (-1)^{\alpha-b} \binom{u-\beta}{b} \binom{u-b}{u-\alpha} \binom{m-u-\beta+b}{b} \sqrt{\binom{m-u-\beta}{v-u}} \binom{v-\beta}{v-u} \binom{v-u+b}{b} \quad \text{for } u \le v \tag{1}$$

and

$$z_{(u,v)}^{\beta\alpha} = \phi_{\beta} z_{\beta\alpha}^{(u,v)} / \left\{ \binom{m}{u} \binom{u}{\alpha} \binom{m-u}{v-u+\alpha} \right\} \quad \text{for } u \le v$$

(see Shirakura and Kuwada, 1976, and Yamamoto *et al.*, 1976), where $\phi_{\beta} = \binom{m}{\beta} - \binom{m}{\beta-1}$. Then some properties of $A_{\beta}^{\#(u,v)}$ are cited in the following:

$$A_{\beta}^{\#(u,w)}A_{\gamma}^{\#(w,v)} = \delta_{\beta\gamma}A_{\beta}^{\#(u,v)},$$

$$\sum_{\beta=0}^{u} A_{\beta}^{\#(u,u)} = I_{\binom{m}{u}}$$
(2)

and

$$\operatorname{rank}\{A_{\beta}^{\#(u,v)}\} = \phi_{\beta},$$

where $\delta_{\beta\gamma}$ is the Kronecker delta.

Let $D_{\alpha}^{(u,v)}(=D_{\alpha}^{(v,u)'})$ $(0 \le \alpha \le u \le v \le \ell)$ and $D_{\beta}^{\#(u,v)}(=D_{\beta}^{\#(v,u)'})$ $(0 \le \beta \le u \le v \le \ell)$ be the matrices of order $v_{\ell}(m)$ such that the (u + 1)-th row block and the (v + 1)-th column block of $D_{\alpha}^{(u,v)}$ and $D_{\beta}^{\#(u,v)}$ are given by $A_{\alpha}^{(u,v)}$ and $A_{\beta}^{\#(u,v)}$, respectively, and zero at elsewhere. Then the information matrix M_T is given by

$$M_T = \sum_{\beta=0}^{\ell} \sum_{u=0}^{\ell-\beta} \sum_{v=0}^{\ell-\beta} \kappa_{\beta}^{u,v} D_{\beta}^{\#(u+\beta,v+\beta)}$$

(see Yamamoto *et al.*, 1976), where T is a BA($N, m, 2, 2\ell; \{\mu_i^{(2\ell)}\}$). Here the relation between $\kappa_{\beta}^{u,v}$ and $\mu_i^{(2\ell)}$ is given by

$$\kappa_{\beta}^{u,v}\left(=\kappa_{\beta}^{v,u}\right) = \sum_{\alpha=0}^{u+\beta} z_{\beta\alpha}^{(u+\beta,v+\beta)} \gamma_{v-u+2\alpha} \quad \text{for } 0 \le u \le v \le \ell - \beta \text{ and } 0 \le \beta \le \ell,$$
(3)

where

$$\gamma_j = \sum_{i=0}^{2\ell} \sum_{p=0}^j (-1)^p {j \choose p} {2\ell - j \choose i - j + p} \mu_i^{(2\ell)} \quad \text{for } 0 \le j \le 2\ell.$$
(4)

The relation between the indices $\mu_i^{(2\ell)}$ of a BA of strength 2ℓ and λ_x of an SA is given by

$$\mu_i^{(2\ell)} = \sum_{x=0}^m \binom{m-2\ell}{x-i} \lambda_x \quad \text{for } 0 \le i \le 2\ell.$$
(5)

Note that size N(=number of assemblies) of an SA $(m; \{\lambda_x\})$ is given by $N = \sum_{x=0}^{m} {m \choose x} \lambda_x$. Furthermore M_T is isomorphic to the symmetric matrices $\|\kappa_{\beta}^{u,v}\| (= K_{\beta}, \text{ say})$ of order $(\ell - \beta + 1)$, i.e., there exists an orthogonal matrix P of order $v_{\ell}(m)$ such that

$$P'M_T P = \text{diag}[K_0; K_1, \cdots, K_1; K_2, \cdots, K_2; \cdots; K_{\ell}, \cdots, K_{\ell}],$$
(6)

where K_{β} ($0 \le \beta \le \ell$) are with multiplicities ϕ_{β} . From (6), the following is immediately:

Lemma 2.1. Let T be an SA(m; $\{\lambda_x\}$). Then the information matrix M_T is non-singular, i.e., T is of resolution $2\ell + 1$, if and only if every K_β ($0 \le \beta \le \ell$) is non-singular, i.e., rank $\{K_\beta\} = \ell - \beta + 1$ for all β .

From (1), and (3) through (5), we have the following (see Hyodo and Yamamoto, 1988):

Lemma 2.2. Let T be an SA(m; { λ_x }). Then we have

$$\kappa_{\beta}^{u,v} = \sum_{x=\beta}^{m-\beta} \left[\left\{ 2^{\beta} \middle/ \sqrt{\binom{m-2\beta}{u}} \right\} \left\{ \sqrt{\lambda_x} \sum_{p=0}^{u} (-1)^p \binom{x-\beta}{u-p} \binom{m-\beta-x}{p} \right\} \sqrt{\binom{m-2\beta}{x-\beta}} \right] \\ \times \left[\left\{ 2^{\beta} \middle/ \sqrt{\binom{m-2\beta}{v}} \right\} \left\{ \sqrt{\lambda_x} \sum_{q=0}^{v} (-1)^q \binom{x-\beta}{v-q} \binom{m-\beta-x}{q} \right\} \sqrt{\binom{m-2\beta}{x-\beta}} \right] \\ for \ 0 \le u \le v \le \ell-\beta, \ 0 \le \beta \le \ell, \ and \ 2\ell \le m.$$

Let F_{β} $(0 \le \beta \le \ell)$ be the $(\ell - \beta + 1) \times (m - 2\beta + 1)$ matrices such that the column vector corresponding to the index λ_x $(x \in V_{\beta})$ is given by $F_{\beta}(x)$, where the (u + 1)-th row of $F_{\beta}(x)$ is given by

$$\sqrt{\lambda_x} \sum_{p=0}^{u} (-1)^p \binom{x-\beta}{u-p} \binom{m-\beta-x}{p} \text{ for } 0 \le u \le \ell - \beta$$
(7)

and $V_{\beta} = \{x \in N_0 | \beta \le x \le m - \beta\}$. Here N_0 is a set of non-negative integers. The (u + 1)-th row and the (v + 1)-th column of K_{β} ($0 \le u, v \le \ell - \beta; 0 \le \beta \le \ell$) correspond to the $(u + \beta)$ -factor interaction and the $(v + \beta)$ -factor one, respectively. Thus the (u + 1)-th row of F_{β} corresponds to the $(u + \beta)$ -factor interaction. Then from (7), we can easily obtain the following theorem (e.g., Hyodo *et al.*, 2015):

Theorem 2.1. Let T be an SA(m; $\{\lambda_x\}$), where $2\ell \leq m$. Then the matrices K_{β} ($0 \leq \beta \leq \ell$) can be expressed as $K_{\beta} = (D_{\beta}F_{\beta}\Lambda_{\beta})(D_{\beta}F_{\beta}\Lambda_{\beta})'$, where D_{β} and Λ_{β} are the diagonal matrices such that the (u + 1)-th element ($0 \leq u \leq \ell - \beta$) of D_{β} and the element of Λ_{β} corresponding to λ_x are given by $2^{\beta}/\sqrt{\binom{m-2\beta}{u}}$ and $\sqrt{\binom{m-2\beta}{x-\beta}}$, respectively.

It follows from Theorem 2.1 that rank{ K_{β} } = r-rank{ F_{β} }, where r-rank{A} denotes the row rank of a matrix A. In order to obtain the rank of a matrix A, we sometimes apply the "elementary row operations" on it. In this case, we positively use the notation "r-rank" instead of the rank. Let $SV_{\beta} = \{x \in V_{\beta} | \lambda_x \neq 0\} (0 \le \beta \le \ell)$, and further let NSV_{β} be the cardinal number of SV_{β} . Then the following is obtained (see Hyodo *et al.*, 2015):

Theorem 2.2. Let *T* be an SA(m; { λ_x }), where $2\ell \le m$. Then it holds that r-rank{ $F_{\beta}(x_1, x_2, \dots, x_{n_{\beta}})$ } = min($n_{\beta}, \ell - \beta + 1$) for { $x_1, x_2, \dots, x_{n_{\beta}}$ } \subset SV $_{\beta}$ ($0 \le \beta \le \ell$), where $F_{\beta}(x_1, x_2, \dots, x_{n_{\beta}})$ = ($F_{\beta}(x_1), F_{\beta}(x_2), \dots, F_{\beta}(x_{n_{\beta}})$). Furthermore the first min($n_{\beta}, \ell - \beta + 1$) rows of $F_{\beta}(x_1, x_2, \dots, x_{n_{\beta}})$ are linearly independent.

The following is due to Ghosh and Kuwada (2001):

Lemma 2.3. Let $K = ||K_{ij}||$ and $L = ||L_{ij}||$ (i, j = 1, 2, 3) be a positive semi-definite matrix of order n with $\operatorname{rank}\{K\} = \operatorname{rank}\left\{\begin{pmatrix}K_{11} & K_{12}\\K_{21} & K_{22}\end{pmatrix}\right\} = n_1 + n_2(\geq 1)$ and some matrix of order n such that $L_{11} = I_{n_1}$ and $L_{1j} = L'_{j1} = O_{n_1 \times n_j}$ (j = 2, 3), respectively, where K_{ij} and L_{ij} are both of size $n_i \times n_j$, and $n_1 + n_2 + n_3 = n$. Then a matrix equation XK = L with parameter matrix X of order n has a solution if and only if

- (i) $n_3 = 0$, where if $n_2 \ge 1$, then L_{22} is arbitrary, or
- (ii) $n_3 \ge 1$, and moreover
 - (1) when $n_2 = 0$, $K_{33} = O_{n_3 \times n_3}$, and furthermore $L_{33} = O_{n_3 \times n_3}$, or
 - (2) when $n_2 \ge 1$, there exists a matrix W of size $n_3 \times n_2$ such that $K_{3j} = WK_{2j}$ (j = 1, 2, 3), and furthermore $L_{i3} = L_{i2}W'$ (i = 2, 3), where L_{i2} are arbitrary.

Remark 2.1. In Lemma 2.3.(i) and (ii)(2), when $n_2 \ge 1$, without loss of generality, we can put $L_{22} = I_{n_2}$ and $L_{23} = L'_{32}$ (= W') (if $n_3 \ge 1$), and hence $L_{33} = WW'$. Furthermore we have $W = K_{32}K_{22}^{-1}$.

Let *T* be an SA(*m*; { λ_x }), where $2\ell \le m$. Then a set of parametric functions *H* Θ is estimable if and only if there exists a matrix *X* of order $v_\ell(m)$ such that $XM_T = H$, where *H* and *X* are given by

$$H = \sum_{\beta=0}^{\ell} \sum_{u=0}^{\ell-\beta} \sum_{v=0}^{\ell-\beta} h_{\beta}^{u,v} D_{\beta}^{\#(u+\beta,v+\beta)} \text{ and } X = \sum_{\beta=0}^{\ell} \sum_{u=0}^{\ell-\beta} \sum_{v=0}^{\ell-\beta} \chi_{\beta}^{u,v} D_{\beta}^{\#(u+\beta,v+\beta)},$$

where $2\ell \le m$. Thus there exist matrices X_{β} such that $X_{\beta}K_{\beta} = H_{\beta}$ for all β ($0 \le \beta \le \ell$) if and only if *T* is of resolution $\mathbb{R}^*(\{1\}|\Omega_{\ell})$, where

$$H_{0} = \begin{pmatrix} h_{0}^{0,0} & 0 & h_{0}^{0,2} & \cdots & h_{0}^{0,\ell} \\ 0 & 1 & 0 & \cdots & 0 \\ h_{0}^{2,0} & 0 & h_{0}^{2,2} & \cdots & h_{0}^{2,\ell} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ h_{0}^{\ell,0} & 0 & h_{0}^{\ell,2} & \cdots & h_{0}^{\ell,\ell} \end{pmatrix}, H_{1} = \begin{pmatrix} 1 & 0 & \cdots & 0 \\ 0 & h_{1}^{1,1} & \cdots & h_{1}^{1,\ell-1} \\ \vdots & \vdots & \ddots & \vdots \\ 0 & h_{1}^{\ell-1,1} & \cdots & h_{1}^{\ell-1,\ell-1} \end{pmatrix}$$

and $H_{\gamma} = \begin{pmatrix} h_{\gamma}^{0,0} & \cdots & h_{\gamma}^{0,\ell-\gamma} \\ \vdots & \ddots & \vdots \\ h_{\gamma}^{\ell-\gamma,0} & \cdots & h_{\gamma}^{\ell-\gamma,\ell-\gamma} \end{pmatrix}$ for $\gamma \ge 2$.

Let $B = \text{diag}[B_1, B_2, B_3]$, $C' = (C'_1 C'_2 C'_3)$ and Δ be a diagonal and non-singular matrix of order $n(=n_1 + n_2 + n_3)$, a matrix of size $n \times p$ with r-rank $\{C\} = \text{r-rank}\left\{\begin{pmatrix} C_1 \\ C_2 \end{pmatrix}\right\} = n_1 + n_2$, and a diagonal and non-singular matrix of order p, respectively, where B_i and C_i (i = 1, 2, 3) are of order n_i and of size $n_i \times p$, respectively. Then we have the following:

Lemma 2.4. Let $Z = ||Z_{ij}||$ (i, j = 1, 2, 3) be a matrix of order $n(= n_1 + n_2 + n_3)$, where Z_{ij} are of size $n_i \times n_j$, and let $K(= ||K_{ij}||) = (BC\Delta)(BC\Delta)'$ and $K_{ij} = (B_iC_i\Delta)(B_jC_j\Delta)'$, and further let $L = ||L_{ij}||$ be a matrix given by Lemma 2.3. Then a matrix equation ZK = L has a solution if and only if

- (i) $n_3 = 0$, where if $n_2 \ge 1$, then L_{22} is arbitrary, or
- (ii) $n_3 \ge 1$, and moreover
 - (1) when $n_2 = 0$, it holds $C_3 = O_{n_3 \times p}$, and furthermore $L_{33} = O_{n_3 \times n_3}$, or
 - (2) when $n_2 \ge 1$, there exists a matrix $W_2^*(=C_3C'_2(C_2C'_2)^{-1})$ of size $n_3 \times n_2$ such that $C_3 = W_2^*C_2$, and furthermore $L_{i3} = L_{i2}(B_3W_2^*B_2^{-1})'$ (i = 2, 3), where L_{i2} are arbitrary.

Proof. (i) When $n_3 = 0$, if $n_2 \ge 1$, then from Lemma 2.3.(i), we have the required result.

- (ii)(1) When $n_3 \ge 1$ and $n_2 = 0$, it follows from Lemma 2.3.(ii)(1) that $K_{33} = (B_3C_3\varDelta)(B_3C_3\varDelta)' = B_3C_3\varDelta\Delta'C'_3B'_3 = 0_{n_3\times n_3}$, and hence $C_3\varDelta\Delta'C'_3 = 0_{n_3\times n_3}$, which implies $C_3\varDelta = 0_{n_3\times p}$. Thus we get $C_3 = 0_{n_3\times p}$, and hence $L_{33} = 0_{n_3\times n_3}$.
- (2) When $n_i \ge 1$ (i = 2, 3), from Lemma 2.3.(ii)(2), there exists a matrix W_2 of size $n_3 \times n_2$ such that $K_{3j} = W_2 K_{2j}$ (j = 1, 2, 3). Since r-rank{C} = r-rank $\left\{ \begin{pmatrix} C_1 \\ C_2 \end{pmatrix} \right\} = n_1 + n_2$, there exists a matrix ($W_1^* \ W_2^*$) of size $n_3 \times (n_1 + n_2)$ such that $C_3 = (W_1^* \ W_2^*) \begin{pmatrix} C_1 \\ C_2 \end{pmatrix}$, where W_k^* (k = 1, 2) are of size $n_3 \times n_k$. Then $K_{3j} = (B_3 C_3 \Delta)(B_j C_j \Delta)' =$ $\{B_3(W_1^* \ W_2^*) \begin{pmatrix} C_1 \\ C_2 \end{pmatrix} \Delta\}(B_j C_j \Delta)' = B_3\{W_1^* B_1^{-1}(B_1 C_1 \Delta)\}(B_j C_j \Delta)' + B_3\{W_2^* B_2^{-1}(B_2 C_2 \Delta)\}(B_j C_j \Delta)' = B_3 W_1^* B_1^{-1} K_{1j} + B_3 W_2^*$ $\times B_2^{-1} K_{2j} = (B_3 W_1^* B_1^{-1} \ B_3 W_2^* B_2^{-1}) \begin{pmatrix} K_{1j} \\ K_{2j} \end{pmatrix}$. Thus we have $(K_{31} \ K_{32}) = (B_3 W_1^* B_1^{-1} \ B_3 W_2^* B_2^{-1}) \begin{pmatrix} K_{11} \ K_{12} \\ K_{21} \ K_{22} \end{pmatrix}$ $= (O_{n_3 \times n_1} \ W_2) \begin{pmatrix} K_{11} \ K_{12} \\ K_{21} \ K_{22} \end{pmatrix}$. Since $\begin{pmatrix} K_{11} \ K_{12} \\ K_{21} \ K_{22} \end{pmatrix}$ is non-singular, we get $B_3 W_1^* B_1^{-1} = O_{n_3 \times n_1}$, i.e., $W_1^* = O_{n_3 \times n_1}$, and $W_2 = B_3 W_2^* B_2^{-1}$.

From Lemma 2.3, the converse is obvious, and hence the required result is obtained.

Let \overline{T} be an array obtained by interchanging all of symbols 0 and 1 of T, where T is an SA(m; $\{\lambda_x\}$). Then it can be easily shown that \overline{T} is also an SA(m; $\{\overline{\lambda}_x\}$), where $\overline{\lambda}_x = \lambda_{m-x}$ for $0 \le x \le m$ (e.g., Shirakura and Kuwada, 1975). Note that \overline{T} is called a complementary SA (CSA) of T. Furthermore if T is of resolution R^{*}($\{1\}|\Omega_\ell$), then \overline{T} is also of resolution R^{*}($\{1\}|\Omega_\ell$).

If $N \ge \nu_{\ell}(m)$, then there always exists a 2^m -BFF design of resolution $2\ell + 1$ (see Hyodo *et al.*, 2015). Thus in the rest of this paper, we consider a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_\ell)$ derived from an SA with $N < \nu_{\ell}(m)$ for $\ell = 2, 3$.

3. Resolution $R^*(\{1\}|\Omega_2)$ designs

We now consider case $\ell = 2$. Then it follows from (7) that $F_{\beta}(x)$ ($0 \le \beta \le 2$) are given by

$$F_0(x) = \sqrt{\lambda_x} \begin{pmatrix} 1 \\ 2x - m \\ \{(2x - m)^2 - m\}/2 \end{pmatrix} \text{ for } x \in V_0,$$
(8a)

$$\boldsymbol{F}_{1}(x) = \sqrt{\lambda_{x}} \begin{pmatrix} 1\\ 2x - m \end{pmatrix} \quad \text{for } x \in V_{1}$$
(8b)

and

$$F_2(x) = \sqrt{\lambda_x(1)} \quad \text{for } x \in V_2. \tag{8c}$$

Let

$$f_{\beta}(x;m) = \{(2x-m)^2 + (2\beta - m)\}/2 \quad \text{for } x \in V_{\beta} \ (\beta = 0, 1).$$
(9)

Then we have the following:

Lemma 3.1. If $f_{\beta}(x_1; m) = f_{\beta}(x_2; m)$ for $\{x_1, x_2\} \subset V_{\beta}$ ($\beta = 0, 1$), then $x_1 + x_2 - m = 0$.

Proof. If $f_{\beta}(x_1;m) = f_{\beta}(x_2;m)$ for $\{x_1, x_2\} \subset V_{\beta}$ ($\beta = 0, 1$), i.e., $\{(2x_1 - m)^2 + (2\beta - m)\}/2 = \{(2x_2 - m)^2 + (2\beta - m)\}/2$, then we have $(2x_1 - m)^2 - (2x_2 - m)^2 = 4(x_1 - x_2)(x_1 + x_2 - m) = 0$. Thus we get the required result.

The following is the main results of this section:

Theorem 3.1. Let T be an SA(m; $\{\lambda_x\}$), where $SV_0 = \{x_1, x_2, \dots, x_{NSV_0}\}$, $N < v_2(m)$ and $m \ge 4$. Then a necessary and sufficient condition for T to be a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_2)$, i.e., of resolution IV, is that non-zero indices of an SA satisfy the following:

- (i) When $NSV_0 = 2$, $x_1 = 1$ and $x_2 = m 1$,
- (ii) when $NSV_0 = 3$,
- (1) $x_1 = 0$, $x_2 = 1$ and $x_3 = m 1$, or its CSA, or

(2)
$$x_1 = 0, x_2 = 2$$
 and $x_3 = 4$, where $m = 4$

and

(iii) when $NSV_0 = 4$, $x_1 = 0$, $x_2 = 1$, $x_3 = m - 1$ and $x_4 = m$.

Proof. See Appendix.

From Theorems 2.1 and 2.2, Lemmas 2.3 and 2.4, and Remark 2.1, we have the following:

Theorem 3.2. If T is a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_2)$, i.e., of resolution IV, derived from an SA($m; \{\lambda_x\}$), where $m \ge 4$ and $N < v_2(m)$, and moreover

(i) when NSV₀ = 2, $A_0^{\#(0,0)}\theta_0 + \left[\left\{1/\sqrt{\binom{m}{2}}\right\}f_0(x_i;m)\right]A_0^{\#(0,2)}\theta_2$ (*i* = 1, 2) and $A_0^{\#(1,1)}\theta_1$ are estimable, where $\{x_1, x_2\} = SV_0$ and $f_0(x;m)$ is given by (9), and furthermore if $f_0(x_i;m) = 0$ for all *i*, then $A_0^{\#(0,0)}\theta_0$ is estimable and $A_0^{\#(2,2)}\theta_2$ is not estimable,

(ii) when NSV₁ = 1, $A_1^{\#(1,1)} \theta_1$ is estimable and $A_1^{\#(2,2)} \theta_2$ is not estimable,

(iii) when NSV₂ = 0, $A_2^{\#(2,2)}\theta_2$ is not estimable

and

(iv) when $\text{NSV}_{\beta} \ge 3 - \beta \ (0 \le \beta \le 2), A_{\beta}^{\#(u,u)} \theta_{u} \ (\beta \le u \le 2)$ are estimable.

Note from (2) that if $A_{\beta}^{\#(u,u)}\theta_u$ are estimable for all β ($0 \le \beta \le u \le 2$), then θ_u is estimable. The results of Theorem 3.1, and estimable parametric functions and the resolution $R(\omega|\Omega_2)$ for each design are summarized in Table 3.1.

Let $K_{\beta}^{(0)}$ $(0 \le \beta < \ell)$ be the matrices of order $(\ell - \beta)$ obtained from K_{β} by cutting off its last row and column, and further let $\mathbf{k}_{\beta}^{1\prime} = (\kappa_{\beta}^{0,0} \kappa_{\beta}^{0,1} \cdots \kappa_{\beta}^{0,\ell-\beta})$ and $\mathbf{k}_{\beta}^{2\prime} = (\kappa_{\beta}^{\ell-\beta,0} \kappa_{\beta}^{\ell-\beta,1} \cdots \kappa_{\beta}^{\ell-\beta,\ell-\beta})$. Then we have the following due to Shirakura (1980):

Proposition 3.1. A necessary and sufficient condition for a BA($N, m, 2, 2\ell; \{\mu_i^{(2\ell)}\}$), T, say, to be a 2^m -BFF design of resolution 2ℓ is that T satisfies the following condition:

For r integers $0 \le \beta_1 < \beta_2 < \cdots < \beta_r \le \ell$ with $|K_{\beta_i}| = 0$ and $|K_{\alpha}| \ne 0$ ($\alpha \ne \beta_j$ ($1 \le j \le r$); $0 \le \alpha \le \ell$),

(i) when $\beta_1 = 0$, there exists a scalar d such that $\mathbf{k}_0^2 = d\mathbf{k}_0^1$, $|\mathbf{K}_0^{(0)}| \neq 0$, $\kappa_{\beta_j}^{\ell-\beta_j,\ell-\beta_j} = 0$ $(1 \le \beta_j \le \ell)$ and $|\mathbf{K}_{\beta_j}^{(0)}| \neq 0$ $(1 \le \beta_j \le \ell-1)$

and

(ii) when $\beta_1 \ge 1$, $\kappa_{\beta_j}^{\ell-\beta_j,\ell-\beta_j} = 0$ $(1 \le \beta_j \le \ell)$ and $|K_{\beta_j}^{(0)}| \ne 0$ $(1 \le \beta_j \le \ell - 1)$,

where |A| denotes the determinant of a matrix A.

In a theoretical sense, Proposition 3.1 above is a very useful result. However it is not always practical. Because the elements $\kappa_{\beta}^{u,v}$ of K_{β} ($0 \le u \le v \le \ell - \beta$; $0 \le \beta \le \ell$) are given by some linear combinations of the indices $\mu_i^{(2\ell)}$ of a BA (or λ_x of an SA) (see (3) and (4)). Hence it is not always easy to obtain $\mu_i^{(2\ell)}$ (or λ_x) such that these indices satisfy some conditions. As an example, we consider case $\ell = 2$:

Let *T* be an SA(*m*; $\{\lambda_x\}$), where $m \ge 4$. Then from (3) through (5), we have

$$\begin{split} \kappa_{0}^{0,0} &= \sum_{x=0}^{m} \binom{m}{x} \lambda_{x} (=N), \\ \kappa_{0}^{0,1} &= \kappa_{0}^{1,0} = \left\{ 1 / \sqrt{\binom{m}{1}} \right\} \sum_{x=0}^{m} (2x-m) \binom{m}{x} \lambda_{x}, \\ \kappa_{0}^{0,2} &= \kappa_{0}^{2,0} = \left[1 / \left\{ 2 \sqrt{\binom{m}{2}} \right\} \right] \sum_{x=0}^{m} \{(2x-m)^{2} - m\} \binom{m}{x} \lambda_{x}, \\ \kappa_{0}^{1,1} &= \left\{ 1 / \binom{m}{1} \right\} \sum_{x=0}^{m} (2x-m)^{2} \binom{m}{x} \lambda_{x}, \\ \kappa_{0}^{1,2} &= \kappa_{0}^{2,1} \right\} = \left[1 / \left\{ 2 \sqrt{\binom{m}{1}\binom{m}{2}} \right\} \right] \sum_{x=0}^{m} (2x-m) \{(2x-m)^{2} - m\} \binom{m}{x} \lambda_{x}, \\ \kappa_{0}^{2,2} &= \left[1 / \left\{ 4\binom{m}{2} \right\} \right\} \sum_{x=0}^{m} \{(2x-m)^{2} - m\}^{2} \binom{m}{x} \lambda_{x}, \\ \kappa_{1}^{0,0} &= 2^{2} \sum_{x=1}^{m-1} \binom{m-2}{x-1} \lambda_{x}, \\ \kappa_{1}^{0,1} &= \kappa_{1}^{1,0} = \left\{ 2^{2} / \sqrt{\binom{m-2}{1}} \right\} \sum_{x=1}^{m-1} (2x-m) \binom{m-2}{x-1} \lambda_{x}, \\ \kappa_{1}^{1,1} &= \left\{ 2^{2} / \binom{m-2}{1} \right\} \sum_{x=1}^{m-1} (2x-m)^{2} \binom{m-2}{x-1} \lambda_{x}. \end{split}$$

and

$$\kappa_2^{0,0} = \sum_{x=2}^{m-2} \binom{m-4}{x-2} \lambda_x.$$

Thus it can be easily shown that $\kappa_1^{1,1} = 0$ if and only if there exists λ_{x^*} such that $(2x^* - m)^2 = 0$ for $x^* \in SV_1$, and $\kappa_2^{0,0} = 0$ if and only if $\lambda_{x^{**}} = 0$ for any $x^{**} \in V_2$. However it is not so easy to obtain the indices λ_x such that $k_0^{2/2}(=(\kappa_0^{2,0} \kappa_0^{2,1} \kappa_0^{2,2})) = dk_0^{1/2}(=d(\kappa_0^{0,0} \kappa_0^{0,1} \kappa_0^{0,2}))$ for $x \in SV_0$ and some d, that is to say, to obtain λ_x such that the system of the linear equations $\kappa_0^{2,u} = d\kappa_0^{0,u}$ (u = 0, 1, 2) satisfies for $x \in SV_0$ and some d. On the other hand, the elements of F_β ($0 \le \beta \le 2$) are given by some polynomial of x of the indices λ_x of an SA (e.g., Hyodo *et al.*, 2015) as seen from (8). In particular, the element of the first row of F_β is all one. Thus it follows from Theorem 2.2 and (8) that if r-rank{ F_0 } = 2 < 3, then there exist two indices λ_{x_i} (i = 1, 2) such that $x_i \in SV_0$ and the elements of the last row of $F_0(x_1, x_2)$ are the same constants, i.e., $f_0(x_1; m) = f_0(x_2; m)$, where $f_0(x; m)$ is given by (9). Next if r-rank{ F_1 } = 1 < 2, then there exists an index λ_{x^*} such that $x^* \in SV_1$ and the last row of $F_1(x^*)$ is 0, i.e., $2x^* - m = 0$, and if r-rank{ F_2 } = 0 < 1, then $F_2(x^{**}) = 0$ for any $x^{**} \in V_2$, i.e., $\lambda_{x^{**}} = 0$. Thus in order to obtain the indices λ_x of an SA such that they satisfy some conditions, the matrices F_β and Theorem 2.2 are very powerful.

Theorem 3.1	3.1 Existence conditions $(\lambda_{x_i} \neq 0 (i = 1, 2, \dots, \text{NSV}_0))$	Estimable parametric functions	Resolution $R(\omega \Omega_2)$
(i)	$x_1 = 1, x_2 = m - 1$	$A_0^{\#(0,0)} \boldsymbol{\theta}_0 + \left[(m-1)(m-4) \left/ \left\{ 2 \sqrt{\binom{m}{2}} \right\} \right] A_0^{\#(0,2)} \boldsymbol{\theta}_2 :$	$\omega = \{0, 1\} \text{ (if } m = 4)$
-		estimable,	$\{1\} (\text{if } m \ge 5)$
-		$A_0^{\#(1,1)}\boldsymbol{\theta}_1$: estimable,	
		$A_0^{\#(2,2)}\theta_2$: not estimable (if $m = 4$),	
-		$A_1^{\#(u,u)}\boldsymbol{\theta}_u$: estimable $(u=1,2)$,	
-		$A_2^{\#(2,2)}\theta_2$: not estimable	
(ii) (1) $x_1 = 0, x_2 = 1, x_3 = m - 1$, or its CSA	$A_{\gamma}^{\#(u,u)}\theta_{u}$: estimable $(\gamma \le u \le 2; \gamma = 0, 1)$,	$\omega = \{0, 1\}$
		$A_2^{\#(2,2)} \theta_2$: not estimable	
(5)	$m=4; x_1=0, x_2=2, x_3=4$	$A_{\gamma}^{\#(u,u)}\theta_{u}$: estimable ($\gamma \leq u \leq 2$; $\gamma = 0, 2$),	
		$A_1^{\#(1,1)} \theta_1$: estimable,	
		$A_1^{\#(2,2)} \theta_2$: not estimable	
(iii)	$x_1 = 0, x_2 = 1, x_3 = m - 1, x_4 = m$	$A_{\gamma}^{\#(u,u)}\theta_{u}$: estimable $(\gamma \leq u \leq 2; \gamma = 0, 1),$	
		$A_2^{\#(2,2)}\theta_2$: not estimable	

Table 3.1. 2^{*m*}-BFF designs of resolution $\mathbb{R}^*(\{1\}|\Omega_2)$ with $N < \nu_2(m)$ ($m \ge 4$)

4. Resolution $R^*(\{1\}|\Omega_3)$ designs

In this section, we consider case $\ell = 3$. By use of the properties of the TMDPB association algebra and the matrix equations, a necessary and sufficient condition for an SA to be a 2^m -BFF design of resolution R^{*}({1}| Ω_3) was already given by Kuwada *et al.* (2003). However their results are very complex. On the other hand, the elements of F_{β} ($0 \le \beta \le 3$) considered here are given by some polynomial of x of the indexes λ_x of an SA as in (10) below, and they are very simple. Thus using these matrices F_{β} and Theorem 2.2, we shall rewrite the existence conditions for a 2^m-BFF design of resolution $R^*(\{1\}|\Omega_3)$ with $N < v_3(m)$.

From (7), $F_{\beta}(x)$ ($0 \le \beta \le 3$) are given by

$$F_0(x) = \sqrt{\lambda_x} \begin{pmatrix} 1 \\ 2x - m \\ \{(2x - m)^2 - m\}/2 \\ (2x - m)\{(2x - m)^2 - (3m - 2)\}/6 \end{pmatrix} \text{ for } x \in V_0,$$
(10a)

$$F_1(x) = \sqrt{\lambda_x} \begin{pmatrix} 1 \\ 2x - m \\ \{(2x - m)^2 - (m - 2)\}/2 \end{pmatrix} \text{ for } x \in V_1,$$
(10b)

$$F_2(x) = \sqrt{\lambda_x} \begin{pmatrix} 1\\ 2x - m \end{pmatrix} \quad \text{for } x \in V_2$$
(10c)

and

$$F_3(x) = \sqrt{\lambda_x} (1) \quad \text{for } x \in V_3. \tag{10d}$$

Let

$$g(x;m) = (2x-m)\{(2x-m)^2 - (3m-2)\}/6 \quad \text{for } x \in V_0 \text{ and } m \ge 6.$$
(11)

Then we have the following:

Lemma 4.1. (I)(i) If $d_{02} = f_0(x_i; m)$ and $d_{03} = g(x_i; m)$ (i = 1, 2) for $\{x_1, x_2\} \subset V_0$ and $m \ge 6$, where d_{0k} (k = 2, 3) are constants, and $f_0(x; m)$ and g(x; m) are given by (9) and (11), respectively, then

(1) when
$$m = 3t^2 + 2t + 1$$
 ($t \ge 1$), we get $x_p = t(3t - 1)/2(\ge 1)$ and $x_q = (t + 1)(3t + 2)/2(\ge 5)$ for $\{p, q\} = \{1, 2\}$ and

and

(2) when
$$m = 3t^2 + 4t + 2$$
 $(t \ge 1)$, we get $x_p = t(3t + 1)/2(\ge 2)$ and $x_q = (t + 1)(3t + 4)/2(\ge 7)$ for $\{p, q\} = \{1, 2\}$.

Here in (1) *and* (2) *just above, we have* $d_{02} = m - 1$ *and* $d_{03} = 0$.

(ii) If $d_0 + f_0(x_i; m)d_2 = g(x_i; m)$ (i = 1, 2, 3) for $\{x_1, x_2, x_3\} \subset V_0$ and $m \ge 6$, where d_k (k = 0, 2) are constants, then $x_p + x_q - m \neq 0 \text{ for some } \{p,q\} \subset \{1,2,3\} \text{ and } (2x_1 - m)(2x_2 - m) + (2x_2 - m)(2x_3 - m) + (2x_3 - m)(2x_1 - m) + (3m - 2) = 0.$ Here

$$d_0 = -\left[(2x_p - m)^2(2x_q - m)^2 - m\{4(x_p - x_q)^2 - (3m - 2)\} - 2(2x_p - m)(2x_q - m)\right]/\{12(x_p + x_q - m)\}$$

and

$$d_2 = \{(2x_p - m)^2 + (2x_p - m)(2x_q - m) + (2x_q - m)^2 - (3m - 2)\}/\{6(x_p + x_q - m)\}.$$

- (II)(i) There does not exist a integer $x_1^* \in V_1$ such that $2x_1^* m = 0$ and $f_1(x_1^*; m) = 0$ for $m \ge 6$, where $f_1(x^*; m)$ is given *by* (9).
- (ii) If $(2x_j^* m)d^* = f_1(x_j^*; m)$ (j = 1, 2) for $\{x_1^*, x_2^*\} \subset V_1$ and $m \ge 6$, where d^* is a constant, then $(2x_1^* m)(2x_2^* m) + (2x_1^* m)(2x_2^* m) + (2x_2^* m) + (2x_2^* m)(2x_2^* m)(2x_2^* m) + (2x_2^* m)(2x_2^* m)(2x_2^* m) + (2x_2^* m)(2x_2^* m)(2x$ (m-2) = 0. Here $d^* = x_1^* + x_2^* - m$.
- *Proof.* (I)(i) If $d_{02} = f_0(x_i;m)$ (i = 1, 2) for $\{x_1, x_2\} \subset V_0$ and $m \ge 6$, i.e., $f_0(x_1;m) = f_0(x_2;m)$, then from Lemma 3.1, we have $x_1 + x_2 - m = 0$. In addition, if $d_{03} = g(x_i; m)$ (i = 1, 2), i.e., $g(x_1; m) = g(x_2; m)$, then $(2x_p - m)\{(2x_p - m)^2 - (2x_p - m)\}$ (3m-2) = 0 for some $p \in \{1, 2\}$. If $2x_p - m = 0$, then $x_p = m/2 = x_q$ for $q \in \{1, 2\} \setminus \{p\}$, and hence $2x_p - m \neq 0$. Thus it must be $(2x_p - m)^2 - (3m - 2) = 0$. Then it has solutions $x_p = (m \pm \sqrt{3m - 2})/2$, which must be integers.

Thus we put $3m - 2 = s^2$ ($s \ge 4$), and hence $m = (s^2 + 2)/3$. If s = 3t ($t \ge 2$), then $m = 3t^2 + 2/3$, if s = 3t + 1 ($t \ge 1$), then $m = 3t^2 + 2t + 1$, and if s = 3t + 2 ($t \ge 1$), then $m = 3t^2 + 4t + 2$, and hence $s \ne 3t$. Since $(2x - m)^2 - (3m - 2) = 0$ for $x \in \{x_1, x_2\}$, we get $d_{02} = m - 1$ and $d_{03} = 0$. Therefore (i) is proved.

- (ii) It follows from Theorem 2.2 that the first three rows of $F_0(x_1, x_2, x_3)$ are linearly independent, and hence the first and the third rows of $F_0(x_1, x_2, x_3)$ are also linearly independent. Thus there exists $\{x_p, x_q\} \subset \{x_1, x_2, x_3\}$ such that $f_0(x_p; m) \neq f_0(x_q; m)$, i.e., $x_p + x_q m \neq 0$. If $d_0 + f_0(x_i; m)d_2 = g(x_i; m)$ (i = 1, 2, 3), i.e., $d_0 + [\{(2x_i m)^2 m\}/2]d_2 = (2x_i m)\{(2x_i m)^2 (3m 2)\}/6$, where $d_k (k = 0, 2)$ are constants, then we get $d_0 = -[(2x_p m)^2(2x_q m)^2 m\{4(x_p x_q)^2 (3m 2)\}/6(x_p + x_q m)]/\{12(x_p + x_q m)\}$ and $d_2 = \{(2x_p m)^2 + (2x_p m)(2x_q m) + (2x_q m)^2 (3m 2)\}/\{6(x_p + x_q m)\}$ for some $\{p, q\} \subset \{1, 2, 3\}$, where $x_p + x_q m \neq 0$. Substituting d_0 and d_2 into $d_0 + f_0(x_r; m)d_2 = g(x_r; m)$ for $r \in \{1, 2, 3\} \setminus \{p, q\}$, we get $(2x_p m)(2x_q m) + (2x_q m)(2x_r m) + (2x_r m)(2x_p m) + (3m 2) = (2x_1 m)(2x_2 m) + (2x_2 m)(2x_3 m) + (2x_3 m)(2x_1 m) + (3m 2) = 0$, and hence (ii) is established.
- (II)(i) If $2x_1^* m = 0$ and $f_1(x_1^*; m) = 0$ for $x_1^* \in V_1$, i.e., $\{(2x_1^* m)^2 (m 2)\}/2 = -(m 2)/2 = 0$, then m = 2 < 6. Thus the required result is obtained.
- (ii) If $(2x_j^* m)d^* = f_1(x_j^*;m)$ (j = 1, 2) for $\{x_1^*, x_2^*\} \subset V_1$ and $m \ge 6$, i.e., $(2x_j^* m)d^* = \{(2x_j^* m)^2 (m 2)\}/2$, then $2(x_1^* x_2^*)d^* = \{(2x_1^* m)^2 (2x_2^* m)^2\}/2 = 2(x_1^* x_2^*)(x_1^* + x_2^* m)$. Thus we get $d^* = x_1^* + x_2^* m$, and hence $(2x_1^* m)(2x_2^* m) + (m 2) = 0$, which is the required result.

The following is the main theorem of this section:

Theorem 4.1. Let T be an SA(m; { λ_x }), where SV₀ = { $x_1, x_2, \dots, x_{NSV_0}$ }, $N < v_3(m)$ and $m \ge 6$. Then a necessary and sufficient condition for T to be a 2^m-BFF design of resolution R^{*}({1}| Ω_3) is that non-zero indices of an SA satisfy the following:

(i) When $NSV_0 = 3$,

(1) $x_1 = 1$, $x_2 = 2$ and $x_3 = 5$, where m = 6, or its CSA,

- (2) $x_1 = 1$, $x_2 = 2$ and $x_3 = 7$, where m = 9, or its CSA, or
- (3) $x_1 = 1$, $x_2 = 3$ and $x_3 = 5$, where m = 6,

(ii) when $NSV_0 = 4$,

(1) x₁ = 0, x₂ = 1, x₃ = 2 and x₄ = m - 1, or its CSA,
(2) x₁ = 0, x₂ = 1, x₃ = m - 2 and x₄ = m - 1, or its CSA,
(3) x₁ = 0, x₂ = 2, x₃ = 4 and x₄ = 6, where m = 6,
(4) x₁ = 0, x₂ = 1, x₃ = 2 and x₄ = m - 2, or its CSA,
(5) x₁ = 0, x₂ = 2, x₃ = m - 2 and x₄ = m - 1, or its CSA,
(6) x₁ = 1, x₂ = 2, x₃ = m - 2 and x₄ = m - 1, where m ≥ 7,
(7) x₁ = 0, x₂ = 1, x₃ = 4 and x₄ = 7, where m = 7, or its CSA,
(8) x₁ = 0, x₂ = 1, x₃ = 3 and x₄ = m - 1, or its CSA,
(9) x₁ = 0, x₂ = 1, x₃ = 4 and x₄ = 7, where m = 8, or its CSA,
(iii) when NSV₀ = 5,

(1) $x_1 = 0, x_2 = 1, x_3 = 2, x_4 = m - 1$ and $x_5 = m$, or its CSA, (2) $x_1 = 0, x_2 = 1, x_3 = 2, x_4 = m - 2$ and $x_5 = m$, or its CSA, (3) $x_1 = 0, x_2 = 1, x_3 = 2, x_4 = m - 2$ and $x_5 = m - 1$, where $m \ge 7$, or its CSA, (4) $x_1 = 0, x_2 = 1, x_3 = 3, x_4 = m - 1$ and $x_5 = m$, or its CSA(if $m \ge 7$), or (5) $x_1 = 0, x_2 = 1, x_3 = 4, x_4 = 7$ and $x_5 = 8$, where m = 8

and

(iv) when $NSV_0 = 6$, $x_1 = 0$, $x_2 = 1$, $x_3 = 2$, $x_4 = m - 2$, $x_5 = m - 1$ and $x_6 = m$, where $m \ge 7$.

Proof. Proof is available in the Appendix.

It follows from Theorems 2.1 and 2.2, Lemmas 2.3, 2.4 and 4.1, and Remark 2.1 that we obtain the following:

Theorem 4.2. If T is a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_3)$ derived from an $SA(m; \{\lambda_x\})$, where $m \ge 6$ and $N < \nu_3(m)$, and furthermore

- (i) when NSV₀ = 3, $A_0^{\#(0,0)}\theta_0 + \left[\left\{1/\sqrt{\binom{m}{3}}\right\}d_0\right]A_0^{\#(0,3)}\theta_3$, $A_0^{\#(1,1)}\theta_1$ and $A_0^{\#(2,2)}\theta_2 + \left[\left\{\sqrt{\binom{m}{2}}/\sqrt{\binom{m}{3}}\right\}d_2\right]A_0^{\#(2,3)}\theta_3$ are estimable, where d_k (k = 0, 2) are given in Lemma 4.1.(I)(ii). In particular, if $d_0 = 0$, then $A_0^{\#(0,0)}\theta_0 = \theta_0$ is estimable and $A_0^{\#(3,3)}\theta_3$ is not estimable, and also if $d_2 = 0$, then $A_0^{\#(2,2)}\theta_2$ is estimable and $A_0^{\#(3,3)}\theta_3$ is not estimable.
- (ii) When NSV₁ = 2, $A_1^{\#(1,1)}\theta_1$ and $A_1^{\#(2,2)}\theta_2 + \left[\left\{\sqrt{\binom{m-2}{1}}/\sqrt{\binom{m-2}{2}}\right\}d^*\right]A_1^{\#(2,3)}\theta_3$ are estimable, where d^* is given in Lemma 4.1.(II)(ii). Particularly if $d^* = 0$, then $A_1^{\#(2,2)}\theta_2$ is estimable and $A_1^{\#(3,3)}\theta_3$ is not estimable.
- (iii) When NSV₂ = 1, $A_2^{\#(2,2)}\theta_2 + \left[\left\{1/\sqrt{\binom{m-4}{1}}\right\}d^{**}\right]A_2^{\#(2,3)}\theta_3$ is estimable, where $d^{**} = 2x^{**} m$ for $x^{**} \in SV_2$. In particular, if $d^{**} = 0$, then $A_2^{\#(2,2)}\theta_2$ is estimable and $A_2^{\#(3,3)}\theta_3$ is not estimable.
- (iv) When NSV₃ = 0, $A_3^{\#(3,3)}\theta_3$ is not estimable.
- (v) When $\text{NSV}_{\beta} \ge 4 \beta \ (0 \le \beta \le 3), A_{\beta}^{\#(u,u)} \theta_u \ (\beta \le u \le 3)$ are estimable.

Analogously to Section 3, if $A_{\beta}^{\#(u,u)}\theta_u$ are estimable for all β ($0 \le \beta \le u \le 3$), then θ_u is estimable. In Table 4.1, the results of Theorem 4.1, and estimable parametric functions and the resolution $R(\omega|\Omega_3)$ for each design are summarized.

5. Discussion

The class of BFF designs is a subset of FF designs. Thus there may exist a better FF design than a BFF design with respect to some criterion (e.g., Kuwada, 1982). However BFF designs possess the same advantage over unbalanced designs as a BIB design does over unbalanced or partially balanced designs. In this paper, we restrict our attention to the class of 2^{*m*}-BFF designs derived from SAs. Under these restrictions, we have given a necessary and sufficient condition for an SA to be a 2^{*m*}-BFF design of resolution R*({1}| Ω_ℓ) for $\ell = 2, 3$, where $N < v_\ell(m)$. As mentioned earlier, if *T* is an SA(*m*; { λ_x }), then it is a BA(*N*, *m*, 2, *t*; { $\mu_i^{(t)}$ }) for any t ($1 \le t \le m$), where the relation between $\mu_i^{(t)}$ and λ_x is given by (5). When m = t + 1, if there exists a BA of strength *t*, then it is an SA. However when m = t + 2, there exists a BA of strength *t* such that it is not always an SA (e.g., Kuriki and Yamamoto, 1984, and Shirakura, 1977). For example,

is a BA($N = 7, m = 4, 2, t = 2; \{\mu_0^{(2)} = 1, \mu_1^{(2)} = \mu_2^{(2)} = 2\}$), but it is not an SA.

Let *T* be an SA(*m*; { $\lambda_1 = \lambda_{m-1} = 1$, $\lambda_x = 0$ ($x \neq 1, m - 1$)}) with N = 2m, where $m \ge 6$. Then under the assumption that the three-factor and higher-order interactions are negligible, $A_0^{\#(0,0)}\theta_0 + \left[(m-1)(m-4)/\left\{2\sqrt{\binom{m}{2}}\right\}\right]A_0^{\#(0,2)}\theta_2$, $A_0^{\#(1,1)}\theta_1$ and $A_1^{\#(u,u)}\theta_u$ (u = 1, 2) are estimable, and $A_2^{\#(2,2)}\theta_2$ is not estimable (see Table 3.1.(i)). Thus from (2), $A_0^{\#(1,1)}\theta_1 + A_1^{\#(1,1)}\theta_1 = \theta_1$ is estimable, and the general mean is confounded with some of the two-factor interaction. On the other hand, under the assumption that the four-factor and higher-order interactions are negligible,

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	Resolution $R(\omega \Omega_3)$	$\omega = \{1\}$									$\omega = \{0, 1, 2\}$					$\omega = \{0, 1\}$				$\omega = \{0, 1, 2\}$		
$(m) \ (m \ge 6)$	Estimable parametric functions	$\begin{array}{l} A_0^{\#(0,0)}\boldsymbol{\theta}_0 + \left(\sqrt{5}/3\right) A_0^{\#(0,3)}\boldsymbol{\theta}_3 : \text{estimable}, \\ A_0^{\#(1,1)}\boldsymbol{\theta}_1 : \text{estimable}, \end{array}$	$A_0^{\#(2,2)} \boldsymbol{ heta}_2 - \left(1/\sqrt{3}\right) A_0^{\#(2,3)} \boldsymbol{ heta}_3$: estimable,	$A_1^{\#(u,u)}\boldsymbol{\theta}_u : \text{ estimable } (u = 1, 2, 3),$ $A_1^{\#(2,2)}\boldsymbol{\theta}_2 - \sqrt{2}A_{\pi^{(2,3)}}^{\#(2,3)}\boldsymbol{\theta}_3 : \text{ estimable,}$	$A_3^{\frac{4}{3}(3,3)}\boldsymbol{\theta}_3$: not estimable	$A_0^{\#(0,0)} \theta_0 + \{28/(3\sqrt{21})A_0^{\#(0,3)} \theta_3 : \text{estimable},$	$A_0^{\#(1,1)}\boldsymbol{\theta}_1$: estimable,	$A_0^{\#(2,2)}\theta_2 - (\sqrt{7/3})A_0^{\#(2,3)}\theta_3$: estimable,	$A_{\gamma}^{\#(u,u)}\theta_u$: estimable ($\gamma \le u \le 3$; $\gamma = 1, 2$),	$A_3^{\#(3,3)} \boldsymbol{\theta}_3$: not estimable	$A_0^{\#(u,u)}\boldsymbol{\theta}_u$: estimable $(u=0, 1, 2),$	$A_0^{\#(3,3)} \theta_3$: not estimable,	$A_{\gamma}^{\#(u,u)}\theta_{u}$: estimable ($\gamma \leq u \leq 3$; $\gamma = 1, 3$),	$A_2^{\#(2,2)} \theta_2$: estimable,	$A_2^{\#(3,3)} \theta_3$: not estimable	$A_{\gamma}^{\#(u,u)}\boldsymbol{\theta}_{u}$: estimable ($\gamma \leq u \leq 3$; $\gamma = 0, 1$),	$A_2^{\#(2,2)}\theta_2 + \{(2x_3 - m)/\sqrt{m - 4}\}A_2^{\#(2,3)}\theta_3:$	estimable,	$A_{\gamma}^{\#(3,3)} \theta_3$: not estimable	$A_{\gamma}^{\#(u,u)}\theta_{u}$: estimable $(\gamma \leq u \leq 3; \gamma = 0, 2),$	$A_1^{\#(u,u)} \boldsymbol{\theta}_u$: estimable $(u=1,2)$,	$A_{\gamma}^{\#(3,3)} \boldsymbol{\theta}_3$: not estimable ($\gamma = 1, 3$)
Table 4.1. 2^m -BFF designs of resolution $\mathbb{R}^*(\{1\} \Omega_3)$ with $N < \nu_3(m)$ $(m \ge 6)$	Existence conditions $(\lambda_{x_i} \neq 0 (i=1, 2, \dots, \text{NSV}_0))$	$m = 6;$ $x_1 = 1, x_2 = 2, x_3 = 5$, or its CSA				$m = 9$; $x_1 = 1$, $x_2 = 2$, $x_3 = 7$, or its CSA					$m = 6; x_1 = 1, x_2 = 3, x_3 = 5$					$x_1 = 0, x_2 = 1, x_3 = 2, x_4 = m - 1$, or its CSA		$x_1 = 0, x_2 = 1, x_3 = m - 2, x_4 = m - 1$, or its CSA		$m=6;$ $x_1=0, x_2=2, x_3=4, x_4=6$		
ole 4.1. 2	Theorem 4.1	(i) (1)				(2)					(3)					(ii) (1)		(2)		(3)		
Та	Th	Ē														(ii)						

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The	orem 4.1	Theorem 4.1 Existence conditions $(\lambda_{x_i} \neq 0 \ (i = 1, 2, \dots, NSV_0))$	Estimable parametric functions	Resolution $R(\omega \Omega_3)$
	(4)	$x_1 = 0, x_2 = 1, x_3 = 2, x_4 = m - 2$, or its CSA	$A_{\gamma}^{\#(u,u)} \theta_u$: estimable ($\gamma \le u \le 3$; $\gamma = 0, 1, 2$), $\omega = \{0, 1, 2\}$	$\omega = \{0, 1, 2\}$
	(5)	$x_1 = 0, x_2 = 2, x_3 = m - 2, x_4 = m - 1$, or its CSA $A_3^{\#(3,3)}\theta_3$: not estimable	$A_3^{\#(3,3)}\theta_3$: not estimable	
	(9)	$m \ge 7$; $x_1 = 1$, $x_2 = 2$, $x_3 = m - 2$, $x_4 = m - 1$		
	(2)	$m = 7$; $x_1 = 0$, $x_2 = 1$, $x_3 = 4$, $x_4 = 7$, or its CSA	$A_{\gamma}^{\#(u,u)} \theta_u$: estimable $(\gamma \le u \le 3; \gamma = 0, 3),$	$\omega = \{0, 1\}$
			$A_1^{\#(1,1)}\boldsymbol{\theta}_1$: estimable,	
			$A_1^{\#(2,2)}\boldsymbol{\theta}_2 - \sqrt{2}A_1^{\#(2,3)}\boldsymbol{\theta}_3$: estimable,	
			$A_2^{\#(2,2)}\boldsymbol{\theta}_2 + \sqrt{1/3}A_2^{\#(2,3)}\boldsymbol{\theta}_3$: estimable	
	(8)	$x_1 = 0, x_2 = 1, x_3 = 3, x_4 = m - 1$, or its CSA	$ A_{\gamma}^{\#(u,u)}\theta_u$: estimable ($\gamma \le u \le 3$; $\gamma = 0, 1, 3$), $ \omega = \{0, 1, 2\}$ (if $m = 6$)	$\omega = \{0, 1, 2\} \text{ (if } m = 6)$
			$\left A_{2}^{\#(2,2)}\boldsymbol{\theta}_{2}+\left\{(2x_{3}-m)/\sqrt{m-4}\right\}A_{2}^{\#(2,3)}\boldsymbol{\theta}_{3}:\right.$	$\{0,1\} (\text{if } m \ge 7)$
	(6)	$m \ge 7$; $x_1 = 0$, $x_2 = 1$, $x_3 = m - 3$, $x_4 = m - 1$, or its CSA	estimable, $\omega = \{0, 1\}$	$\omega = \{0, 1\}$
			$A_2^{\#(3,3)}\theta_3$: not estimable (if $2x_3 - m = 0$)	
	(10)	$m = 8$; $x_1 = 0$, $x_2 = 1$, $x_3 = 4$, $x_4 = 7$, or its CSA		$\omega = \{0, 1, 2\}$
(iii) (1)	(1)	$x_1 = 0, x_2 = 1, x_3 = 2, x_4 = m - 1, x_5 = m,$	$A_{\gamma}^{\#(u,u)}\boldsymbol{\theta}_{u}$: estimable ($\gamma \leq u \leq 3$; $\gamma = 0, 1$),	$\omega = \{0, 1\}$
		or its CSA	$A_{2}^{\#(2,2)}\boldsymbol{\theta}_{2} + \left\{ (2x_{3} - m)/\sqrt{m - 4} \right\} A_{2}^{\#(2,3)}\boldsymbol{\theta}_{3}:$	
			estimable,	
			$A_3^{\#(3,3)} \theta_3$: not estimable	
	(2)	$x_1 = 0, x_2 = 1, x_3 = 2, x_4 = m - 2, x_5 = m,$	$\left A_{\gamma}^{\#(u,u)}\boldsymbol{\theta}_{u}:$ estimable $(\gamma \leq u \leq 3; \gamma = 0, 1, 2), \omega = \{0, 1, 2\}$	$\omega = \{0, 1, 2\}$
		or its CSA	$A_3^{\#(3,3)}\theta_3$: not estimable	
	(3)	$m \ge 7$; $x_1 = 0$, $x_2 = 1$, $x_3 = 2$, $x_4 = m - 2$, $x_5 = m - 1$,		
		or its CSA		
	(4)	$x_1 = 0, x_2 = 1, x_3 = 3, x_4 = m - 1, x_5 = m,$	$ A_{\gamma}^{\#(u,u)}\theta_u$: estimable ($\gamma \le u \le 3$; $\gamma = 0, 1, 3$), $ \omega = \{0, 1, 2\}$ (if $m = 6$)	$\omega = \{0, 1, 2\} \text{ (if } m = 6)$
		or its CSA (if $m \ge 7$)	$A_{2}^{\#(2,2)}\theta_{2} + \{(2x_{3}-m)/\sqrt{m-4}\}A_{2}^{\#(2,3)}\theta_{3}:$	$\{0, 1\}$ (if $m \ge 7$)
			estimable,	
			$A_2^{\#(3,3)}\theta_3$: not estimable (if $2x_3 - m = 0$)	
	(5)	$m = 8; x_1 = 0, x_2 = 1, x_3 = 4, x_4 = 7, x_5 = 8$		$\omega = \{0, 1, 2\}$
(iv)		$m \ge 7$; $x_1 = 0$, $x_2 = 1$, $x_3 = 2$, $x_4 = m - 2$, $x_5 = m - 1$,	$A_{\gamma}^{\#(u,u)}\theta_u$: estimable ($\gamma \le u \le 3$; $\gamma = 0, 1, 2$),	
		$x^0 = m$	$A_3^{\#(3,3)}\theta_3$: not estimable	

Table 4.1. (Continued)

$$\begin{aligned} \text{r-rank} &\{F_0(1, m-1)\} \\ &= \text{r-rank} \left\{ \begin{pmatrix} 1 & 1 \\ -(m-2) & m-2 \\ (m-1)(m-4)/2 & (m-1)(m-4)/2 \\ -(m-1)(m-2)(m-6)/6 & (m-1)(m-2)(m-6)/6 \end{pmatrix} \right\} \\ &= 2 < 4, \\ \text{r-rank} &\{F_1(1, m-1)\} \\ &= \text{r-rank} \left\{ \begin{pmatrix} 1 & 1 \\ -(m-2) & m-2 \\ (m-2)(m-3)/2 & (m-2)(m-3)/2 \end{pmatrix} \right\} \\ &= 2 < 3 \end{aligned}$$

and

$$r-rank{F_2} = r-rank{F_3} = 0$$

Thus the third row of $F_0(1, m-1)$ equals (m-1)(m-4)/2 times the first and its last row equals (m-1)(m-6)/6 times the second, and the last row of $F_1(1, m-1)$ equals (m-2)(m-3)/2 times the first. Hence from Lemma 2.4, $A_0^{\#(0,0)}\theta_0 + \{(m-4)\sqrt{(m-1)/(2m)}\}A_0^{\#(0,2)}\theta_2, A_0^{\#(1,1)}\theta_1 + [(m-6)\sqrt{(m-1)/(6(m-2))}]A_0^{\#(1,3)}\theta_3, A_1^{\#(1,1)}\theta_1 + \{\sqrt{\binom{m-2}{2}}\}A_1^{\#(1,3)}\theta_3$ and $A_1^{\#(2,2)}\theta_2$ are estimable, and $A_{\gamma}^{\#(u,u)}\theta_u$ ($2 \le \gamma \le u \le 3$) are not estimable. This implies that the main effect is confounded with some of the three-factor interaction. Therefore if the three-factor and higher-order interactions are negligible, then *T* is of resolution $\mathbb{R}^*(\{1\}|\Omega_2)$, and hence it is also of resolution $\mathbb{R}^*(\{1\}|\Omega_3)$.

Appendix

In this Appendix, we provide the proofs of Theorems 3.1 and 4.1.

Proof of Theorem 3.1. We shall prove the claim by listing out all possible cases. In these cases, since $SV_0 \supset SV_1 \supset SV_2$, we have $NSV_0 \ge NSV_1 \ge NSV_2$, and $NSV_{\gamma} - NSV_{\gamma+1} \le 2$ for $\gamma = 0, 1$. Furthermore when $NSV_0 \ge 5$, we have $N \ge 1 + 1 + m + m + {m \choose 2} > v_2(m)$ for $m \ge 4$, and hence $NSV_0 \le 4$. Then the proof starts with case $NSV_2 = 0$.

- [A] When NSV₂ = 0, i.e., $\lambda_{x^{**}} = 0$ for any $x^{**} \in V_2$, from Theorem 2.2, it must be that $1 \leq \text{NSV}_1 \leq 2$ and $\text{NSV}_0 \geq 2$. In addition,
- [a] when NSV₁ = 1, i.e., $x_1^* = 1$ or m 1, we have r-rank{ $F_1(x_1^*)$ } = 1 < 2. Thus from (8b), the last row of $F_1(x_1^*)$ must be 0, i.e., $2x_1^* m = 0$, where $m = 2s \ge 4$. However $2x_1^* m = -(m 2) < 0$ and m 2 > 0 for $m \ge 4$ according as $x_1^* = 1$ and m 1, respectively. Therefore in this case, there does not exist a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_2)$.
- [b] When NSV₁ = 2, i.e., $x_1^* = 1$ and $x_2^* = m 1$, we have r-rank{ $F_1(x_1^*, x_2^*)$ } = 2.
 - [1] When NSV₀=2, i.e., $x_i = x_i^*$ (i = 1, 2), we have r-rank{ $F_0(x_1, x_2)$ } = 2 < 3. Thus from (8a), the elements of the last row of $F_0(x_1, x_2)$ must be the same, i.e., $f_0(x_1; m) = f_0(x_2; m)$, where $f_0(x; m)$ is given by (9). Since $x_1 + x_2 m = 0$ and $v_2(m) N \le v_2(m) 2m = (m 4)(m + 1)/2 + 3 > 0$ for $m \ge 4$, it follows from Lemma 3.1 that case (i) is established.
 - [2] When NSV₀ = 2 + p (p = 1, 2), i.e., x_1 = 0 or m (if p = 1) (x_1 = 0 (if p = 2)) and $x_{i+1} = x_i^*$ (i = 1, 2) (and $x_4 = m$ (if p = 2)), we have r-rank{ $F_0(x_1, x_2, \dots, x_{2+p})$ } = 3. Then $v_2(m) N \le (m-4)(m+1)/2 + (3-p) > 0$ for $m \ge 4$ and p = 1, 2. Thus cases (ii)(1) and (iii) are proved.
- [B] When NSV₂ = 1, i.e., $2 \le x_1^{**} \le m-2$, we have r-rank{ $F_2(x_1^{**})$ } = 1. When NSV₁ = 3, we have $N \ge 2m + \binom{m}{2} > v_2(m)$ for $m \ge 4$, and hence (NSV₂ = 1 \le)NSV₁ ≤ 2 . Furthermore from Theorem 2.2, it must be NSV₀ ≥ 2 .
- [a] When NSV₁ = 1, i.e., $x_1^* = x_1^{**}$, we have r-rank{ $F_1(x_1^*)$ } = 1 < 2. Thus from (8b), it must be $2x_1^* m = 0$, and hence $x_1^* = m/2$, where $m = 2s \ge 4$. Furthermore
 - [1] when NSV₀ = 2, i.e., $x_1 = 0$ or m and $x_2 = x_1^*$, we have r-rank{ $F_0(x_1, x_2)$ } = 2 < 3. Thus it must be $f_0(x_1; m) = f_0(x_2; m)$ for $m \ge 4$. However $x_1 + x_2 m \ne 0$ for $m \ge 4$, and hence, from Lemma 3.1, there does not exist a 2^m -BFF design of resolution R^{*}({1}|\Omega_2).

- [2] When NSV₀ = 3, i.e., $x_1 = 0$, $x_2 = x_1^*$ and $x_3 = m$, we have r-rank{ $F_0(x_1, x_2, x_3)$ } = 3. When m = 4, we have $8 \le N < v_2(4) = 11$, and when $m = 2s \ge 6$, $v_2(m) N \le v_2(m) \{2 + \binom{m}{s}\} \le \binom{m}{2} + m \binom{m}{3} 1 = -(m-6)\{(m-6)(m+6) + 35\}/6 \le 0$ for $m = 2s \ge 6$. Thus we get m = 4, and hence $x_1 = 0$, $x_2 = m/2 = 2$ and $x_3 = 4$, which is case (ii)(2).
- [b] When NSV₁ = 2, i.e., $x_1^* = 1$ or m 1 and $x_2^* = x_1^{**}$, we have r-rank{ $F_1(x_1^*, x_2^*)$ } = 2. In this case, $N \ge m + \binom{m}{2} = v_2(m) 1$, and hence it must be NSV₀ = 2, i.e., $x_i = x_i^*$ (i = 1, 2). Then we have r-rank{ $F_0(x_1, x_2)$ } = 2 < 3. Thus from (8a), it must be $f_0(x_1; m) = f_0(x_2; m)$ for $m \ge 4$. However $x_1 + x_2 m \ne 0$ for $m \ge 4$, and hence, from Lemma 3.1, a 2^m-BFF design of resolution R^{*}({1}|\Omega_2) does not exist.
- [C] When NSV₂ = $q \ge 2$, where $(m 2) 1 \ge q$, we have $N \ge q\binom{m}{2} \ge 2\binom{m}{2} > v_2(m)$ for $m \ge 3 + q$. Thus in this case, there does not exist a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_2)$ with $N < v_2(m)$ for $m \ge 3 + q$.

Proof of Theorem 4.1. Similarly to the proof of Theorem 3.1, it will be done by listing out all possible cases. In these cases, we have $SV_0 \supset SV_1 \supset SV_2 \supset SV_3$, and hence $NSV_0 \ge NSV_1 \ge NSV_2 \ge NSV_3$, and $NSV_{\gamma} - NSV_{\gamma+1} \le 2$ for $\gamma = 0, 1, 2$. Moreover when $NSV_0 \ge 7$, $N \ge 1 + 1 + m + m + {m \choose 2} + {m \choose 3} > v_3(m)$ for $m \ge 6$, and hence $NSV_0 \le 6$. We also begin the proof with $NSV_3 = 0$.

- [A] When NSV₃ = 0 (and hence NSV₂ \leq 2), i.e., $\lambda_{x^{***}} = 0$ for any $x^{***} \in V_3$, it follows from Theorem 2.2 that NSV₁ \geq 1 and NSV₀ \geq 2. Moreover
 - [a] when $NSV_2 = 0$ (and hence $NSV_1 \le 2$), and furthermore
 - [1] when NSV₁ = 1, i.e., $x_1^* = 1$ or m 1, we have r-rank{ $F_1(x_1^*)$ } = 1 < 3. Thus from Lemma 4.1.(II)(i), there does not exist a 2^m -BFF design of resolution R^{*}({1}|\Omega_3).
 - [2] When NSV₁ = 2, i.e., $x_1^* = 1$ and $x_2^* = m 1$, r-rank{ $F_1(x_1^*, x_2^*)$ } = 2 < 3. Then we have $(2x_1^* m)(2x_2^* m) + (m 2) = -(m 2)(m 3) < 0$ for $m \ge 6$. Therefore from Lemma 4.1.(II)(ii), a 2^m-BFF design of resolution R^{*}({1}|\Omega_3) does not exist.
 - [b] When NSV₂ = 1 (and hence NSV₁ \leq 3), i.e., $x_1^{**} = 2$ or m 2, we have r-rank{ $F_2(x_1^{**})$ } = 1 < 2.
 - [1] When NSV₁ = 1, i.e., $x_1^* = x_1^{**}$, we have r-rank{ $F_1(x_1^*)$ } = 1 < 3. Then from Lemma 4.1.(II)(i), there does not exist a 2^m -BFF design of resolution R^{*}({1}| Ω_3).
 - [2] When NSV₁ = 2, i.e., $x_1^* = 1$ or m 1 and $x_2^* = x_1^{**}$, we have r-rank{ $F_1(x_1^*, x_2^*)$ } = 2 < 3. When $x_1^* = 1$, $(2x_1^* m)(2x_2^* m) + (m 2) = (m 2)(m 3) > 0$ and -(m 2)(m 5) < 0 for $m \ge 6$ according as $x_2^* = 2$ and m 2, respectively. Thus, it follows from Lemma 4.1.(II)(ii) and the relation of the CSA that there does not exist a 2^m -BFF design of resolution R^{*}({1}|\Omega_3).
 - [3] When NSV₁ = 3 (and hence NSV₀ \leq 5), i.e., $x_1^* = 1$, $x_2^* = x_1^{**}$ and $x_3^* = m 1$, r-rank{ $F_1(x_1^*, x_2^*, x_3^*)$ } = 3. In addition,
 - [3.1] when NSV₀ = 3, i.e., $x_i = x_i^*$ (*i* = 1, 2, 3), r-rank{ $F_0(x_1, x_2, x_3)$ } = 3 < 4. Then $x_1 + x_2 m \neq 0$ for $x_2 = 2$ or m 2, and $m \ge 6$. Furthermore from Lemma 4.1.(I)(ii), we have $(2x_1 m)(2x_2 m) + (2x_2 m)(2x_3 m) + (2x_3 m)(2x_1 m) + (3m 2) = -(m 1)(m 6) = 0$ for $x_2 = 2$ or m 2. Thus we get m = 6, and hence $x_3 = m 1 = 5$. In this case, $27 \le N < v_3(6) = 42$. Therefore case (i)(1) is established. Here if $x_2 = 2$, then $d_0 = 10/3$ and $d_2 = -2/3$, and if $x_2 = m 2 = 4$, then $d_0 = -10/3$ and $d_2 = 2/3$, where d_k (k = 0, 2) are constants given in Lemma 4.1.(I)(ii).
 - [3.2] When NSV₀ = 3 + p (p = 1, 2), i.e., $x_1 = 0$ or m (if p = 1) ($x_1 = 0$ (if p = 2)), $x_2 = 1$, $x_3 = 2$ or m 2, and $x_4 = m 1$ (and $x_5 = m$ (if p = 2)), we have r-rank{ $F_0(x_1, x_2, \dots, x_{3+p})$ } = 4. Then $\binom{m}{2} + 2m + p \le N < v_3(m)$ for $m \ge 6$ and p = 1, 2. Thus we obtain cases (ii)(1) and (2), and (iii)(1).
 - [c] When NSV₂ = 2 (and hence NSV₁ ≤ 4), i.e., $x_1^{**} = 2$ and $x_2^{**} = m 2$, r-rank{ $F_2(x_1^{**}, x_2^{**})$ } = 2.
 - [1] When NSV₁ = 2 (and hence NSV₀ ≤ 4), i.e., $x_j^* = x_j^{**}$ (j = 1, 2), r-rank{ $F_1(x_1^*, x_2^*)$ } = 2 < 3. Thus from Lemma 4.1.(II)(ii), $(2x_1^* m)(2x_2^* m) + (m 2) = -(m 3)(m 6) = 0$ for $m \ge 6$, and hence we get m = 6 and $x_2^* = m 2 = 4$. In this case, $d^* = 0$, where d^* is a constant given in Lemma 4.1.(II)(ii). Furthermore
 - [1.1] when NSV₀ = 2, i.e., $x_i = x_i^*$ (i = 1, 2), r-rank{ $F_0(x_1, x_2)$ } = 2 < 4. Since m = 6, from Lemma 4.1.(I)(i)(1), we get t = 1, and hence $x_1 = 1$ and $x_2 = 5$, which contradict $x_1 = 2$ and $x_2 = 4$. Thus in this case, there does not exist a 2^m -BFF design of resolution R^{*}({1}|\Omega_3).

- [1.2] When NSV₀ = 3, i.e., $x_1 = 0$ or m(= 6) and $x_{i+1} = x_i^*$ (i = 1, 2), r-rank{ $F_0(x_1, x_2, x_3)$ } = 3 < 4. Since m = 6, we have $x_1 + x_2 m = -4 \neq 0$ and $(2x_1 m)(2x_2 m) + (2x_2 m)(2x_3 m) + (2x_3 m)(2x_1 m) + (3m 2) = 12 \neq 0$ for $x_1 = 0$. Therefore from Lemma 4.1.(I)(ii) and the relation of the CSA, a 2^m -BFF design of resolution R*({1}|\Omega_3) does not exist.
- [1.3] When NSV₀ = 4, i.e., $x_1 = 0$, $x_{i+1} = x_i^*$ (i = 1, 2) and $x_4 = m(= 6)$, r-rank{ $F_0(x_1, x_2, \dots, x_4)$ } = 4. In this case, $32 \le N < v_3(6) = 42$, and hence case (ii)(3) is proved.
- [2] When NSV₁ = 3 (and hence NSV₀ \leq 5), i.e., $x_1^* = 1$ or m 1 and $x_{j+1}^* = x_j^{**}$ (j = 1, 2), we have r-rank{ $F_1(x_1^*, x_2^*, x_3^*)$ } =3.
 - [2.1] When NSV₀ = 3, i.e., $x_i = x_i^*$ (i = 1, 2, 3), r-rank{ $F_0(x_1, x_2, x_3)$ } = 3 < 4. Then $x_1 + x_2 m = -(m-3)(\neq 0)$ and $1(\neq 0)$ for $m \ge 6$ according as $x_1 = 1$ and m - 1, respectively. Furthermore $(2x_1 - m)(2x_2 - m) + (2x_2 - m)(2x_3 - m) + (2x_3 - m)(2x_1 - m) + (3m - 2) = -(m - 2)(m - 9)$ for $x_1 = 1$ or m - 1. Thus from Lemma 4.1.(I)(ii), we get m = 9, and hence $x_1 = 1$ or m - 1 = 8 and $x_3 = m - 2 = 7$. In this case, $81 \le N < v_3(9) = 130$. Thus case (i)(2) is established. If $x_1 = 1$, $x_2 = 2$ and $x_3 = 7$, then we get $d_0 = 56/3$ and $d_2 = -7/3$, and if $x_1 = 8$, $x_2 = 2$ and $x_3 = 7$, then $d_0 = -56/3$ and $d_2 = 7/3$.
 - [2.2] When NSV₀ = 3 + p (p = 1, 2), i.e., $x_1 = 0$ or m (if p = 1) ($x_1 = 0$ (if p = 2)) and $x_{i+1} = x_i^*$ (i = 1, 2, 3)(and $x_5 = m$ (if p = 2)), we have r-rank{ $F_0(x_1, x_2, \dots, x_{3+p})$ } = 4. Furthermore $v_3(m) N \le (m-6)(m^2+5)/6 + (6-p) > 0$ for $m \ge 6$ and p = 1, 2. Thus cases (ii)(4) and (5), and (iii)(2) are proved.
- [3] When NSV₁ = 4, i.e., $x_1^* = 1$, $x_2^* = 2$, $x_3^* = m 2$ and $x_4^* = m 1$, r-rank{ $F_1(x_1^*, x_2^*, \dots, x_4^*)$ } = 3. Furthermore when NSV₀ = 4 + p (p = 0, 1, 2), i.e., when p = 0, we have $x_i = x_i^*$ ($i = 1, 2, \dots, 4$), when p = 1, $x_1 = 0$ or m and $x_{i+1} = x_i^*$ ($i = 1, 2, \dots, 4$), and when p = 2, $x_1 = 0$, $x_{i+1} = x_i^*$ ($i = 1, 2, \dots, 4$) and $x_6 = m$, we have r-rank{ $F_0(x_1, x_2, \dots, x_{4+p})$ } = 4. Then $v_3(m) N \le (m 6)(m^2 1)/6 p = (m 7)(m^2 + m + 6)/6 + (8 p)$ for $m \ge 6$ and p = 0, 1, 2. Thus when m = 6, we have $v_3(6) N \le 0$ for p = 0, 1, 2, and when $m \ge 7$, $v_3(m) N > 0$ for p = 0, 1, 2. Therefore we establish cases (ii)(6), (iii)(3) and (iv).
- [B] When NSV₃ = 1 (and hence NSV₂ \leq 3), i.e., $3 \leq x_1^{***} \leq m-3$, it follows from Theorem 2.2 that r-rank{ $F_3(x_1^{***})$ } = 1, and it must be NSV₀ \geq 2. In addition,
 - [a] when NSV₂ = 1 (and hence NSV₁ \leq 3), i.e., $x_1^{**} = x_1^{***}$, we have r-rank{ $F_2(x_1^{**})$ } = 1 < 2.
 - [1] When NSV₁= 1, i.e., $x_1^* = x_1^{**}$, r-rank{ $F_1(x_1^*)$ } = 1 < 3. Thus from Lemma 4.1.(II)(i), there does not exist a 2^m -BFF design of resolution R^{*}({1}|\Omega_3).
 - [2] When NSV₁ = 2 (and hence NSV₀ ≤ 4), i.e., $x_1^* = 1$ or m 1 and $x_2^* = x_1^{**}$, r-rank{ $F_1(x_1^*, x_2^*)$ } = 2 < 3. Then from Lemma 4.1.(II)(ii), it must be $(2x_1^* m)(2x_2^* m) + (m 2) = 0$ for $m \ge 6$. We consider case $x_1^* = 1$. Then we have $(2x_1^* m)(2x_2^* m) + (m 2) = -(m 2)\{2x_2^* (m + 1)\} = 0$ for $m \ge 6$, and hence $x_2^* = (m + 1)/2$, where $m = 2s + 1 \ge 7$. Moreover
 - [2.1] when NSV₀ = 2, i.e., $x_i = x_i^*$ (i = 1, 2), we have r-rank{ $F_0(x_1, x_2)$ } = 2 < 4. Since $x_1 = 1$, from Lemma 4.1.(I)(i)(1), we get t = 1, and hence m = 6 and $x_2 = 5$. However m = 6 < 7 and $5 \notin SV_3$ for m = 6. Thus from Lemma 4.1.(I)(i)(1) and the relation of the CSA, a 2^m-BFF design of resolution R^{*}({1}|\Omega_3) does not exist.
 - [2.2] When NSV₀ = 3, i.e., $x_1 = 0$ or m and $x_{i+1} = x_i^*$ (i = 1, 2), we have r-rank{ $F_0(x_1, x_2, x_3)$ } = 3 < 4. Then it holds $x_p + x_q m \neq 0$ for any {p, q} \subset {1, 2, 3} and $m \geq 7$. However when $x_2 = 1$, $(2x_1 m)(2x_2 m) + (2x_2 m)(2x_3 m) + (2x_3 m)(2x_1 m) + (3m 2) = m(m 1) > 0$ and -m(m 5) < 0 for $m \geq 7$ according as $x_1 = 0$ and m, respectively. Therefore it follows from Lemma 4.1.(I)(ii) and the relation of the CSA that there does not exist a 2^m -BFF design of resolution R*({1}|\Omega_3).
 - [2.3] When NSV₀ = 4, i.e., $x_1 = 0$ and $x_{i+1} = x_i^*$ (i = 1, 2) and $x_4 = m$, r-rank{ $F_0(x_1, x_2, \dots, x_4)$ } = 4. In this case, when m = 7, we have $44 \le N < v_3(7) = 64$, and when $m = 2s + 1 \ge 9$, $v_3(m) - N \le \binom{m}{2} + \binom{m}{3} - \binom{m}{4} - 1 < \binom{m}{2} + \binom{m}{3} - \binom{m}{4} = -[(m-9)\{(m-9)(m^2 + 8m + 74) + 682\}/24 + 6] < 0$ for $m \ge 9$. Thus we get m = 7, and hence $x_1 = 0$, $x_2 = 1$ or m - 1 = 6, $x_3 = (m+1)/2 = 4$ and $x_4 = m = 7$. Therefore case (ii)(7) is proved.
 - [3] When NSV₁ = 3 (and hence NSV₀ \leq 5), i.e., $x_1^* = 1$, $x_2^* = x_1^{**}$ and $x_3^* = m 1$, we have r-rank{ $F_1(x_1^*, x_2^*, x_3^*)$ } = 3.
 - [3.1] When NSV₀=3, i.e., $x_i = x_i^*$ (i = 1, 2, 3), r-rank{ $F_0(x_1, x_2, x_3)$ } = 3 < 4. Then $x_1 + x_2 m < 0$ for $m \ge 6$. Furthermore from Lemma 4.1.(I)(ii), it must be $(2x_1 - m)(2x_2 - m) + (2x_2 - m)(2x_3 - m) + (2x_3 - m)(2x_1 - m) + (3m - 2) = -(m - 1)(m - 6) = 0$ for $m \ge 6$. Thus we get m = 6, and hence, $x_1 = 1, x_2 = 3$ and $x_3 = m - 1 = 5$. In this case, $32 \le N < v_3(6) = 42$, and hence we have case (i)(3), and $d_0 = d_2 = 0$.

- [3.2] When NSV₀ = 3 + p (p = 1, 2), i.e., $x_1 = 0$ or m (if p = 1) ($x_1 = 0$ (if p = 2)), $x_{i+1} = x_i^*$ (i = 1, 2) and $x_4 = m 1$ (and $x_5 = m$ (if p = 2)), r-rank{ $F_0(x_1, x_2, \dots, x_{3+p})$ } = 4. When $x_3 = 3$ or m 3 (if $m \ge 7$), $v_3(m) N \le (m 6)(m + 3)/2 + (10 p) > 0$ for $m \ge 6$ and p = 1, 2, and hence we obtain cases (ii)(8) and (9), and (iii)(4). When $x_3 = 4$ and $m = 8, v_3(8) N \le 7 p > 0$ for p = 1, 2. Thus cases (ii)(10) and (iii)(5) are obtained. Furthermore when $4 \le x_3 \le m 4$ and $m \ge 9$, then $N \ge p + 2m + {m \choose x_3} \ge p + 2m + {m \choose 4}$, and hence $v_3(m) N \le {m \choose 2} + {m \choose 3} m {m \choose 4} (p 1) < {m \choose 2} + {m \choose 3} {m \choose 4} < 0$ for $m \ge 9$ and p = 1, 2 (see [B][a][2][2.3] above). Thus there does not exist a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_3)$ with $N < v_3(m)$ for $m \ge 9$.
- [b] When NSV₂=2 (and hence NSV₁ ≤ 4), i.e., $x_1^{**} = 2$ or m 2 and $x_2^{**} = x_1^{***}$, we have r-rank{ $F_2(x_1^{**}, x_2^{**})$ } = 2.
 - [1] When NSV₁=2, i.e., $x_j^* = x_j^{**}$ (j = 1, 2), we have r-rank{ $F_1(x_1^*, x_2^*)$ } = 2 < 3. Thus from Lemma 4.1.(II)(ii), it must be $(2x_1^* m)(2x_2^* m) + (m 2) = 0$ for $m \ge 6$. When $x_1^* = 2$, we have $x_2^* = (m + 1)/2 + 1/(m 4)$, which is an integer for m = 6 only, and hence $x_2^* = 4$. However 4 does not belong to SV_3 for m = 6. Therefore from Lemma 4.1.(II)(ii) and the relation of the CSA, there does not exist a 2^m -BFF design of resolution R*({1}| Ω_3).
 - [2] When NSV₁ = 3 (and hence NSV₀ \leq 5), i.e., $x_1^* = 1$ or m 1 and $x_{j+1}^* = x_j^{**}(j = 1, 2)$, we have r-rank{ $F_1(x_1^*, x_2^*, x_3^*)$ } = 3. In this case, $N \geq m + {m \choose 2} + {m \choose 2}$ for $m \geq 6$. Thus when $4 \leq x_3^* \leq m 4$ and $m \geq 8$, $N \geq m + {m \choose 2} + {m \choose 4} > v_3(m)$, and hence we only consider case $x_3^* = 3$ or m 3 (if $m \geq 7$) for $m \geq 6$.
 - [2.1] When NSV₀ = 3, i.e., $x_i = x_i^*$ (i = 1, 2, 3), we have r-rank{ $F_0(x_1, x_2, x_3)$ } = 3 < 4. Then $x_s + x_t m \neq 0$ for any {s, t} \subset {1, 2, 3}. Moreover when $x_1 = 1$, $(2x_1 - m)(2x_2 - m) + (2x_2 - m)(2x_3 - m) + (2x_3 - m)(2x_1 - m) + (3m - 2) = 3(m^2 - 7m + 14), -(m^2 - 15m + 30), -(m^2 - 11m + 22)$ and $-(m^2 - 7m - 2)$ according as $x_2 = 2$ and $x_3 = 3$, $x_2 = 2$ and $x_3 = m - 3$, $x_2 = m - 2$ and $x_3 = 3$, and $x_2 = m - 2$ and $x_3 = m - 3$, respectively. However $m^2 - 7m + 14 = (m - 6)(m - 1) + 8 > 0$ for $m \ge 6$, and furthermore three quadratic equations $m^2 - 15m + 30 = 0$, $m^2 - 11m + 22 = 0$ and $m^2 - 7m - 2 = 0$ do not have an integer solution for $m \ge 6$. Therefore from Lemma 4.1.(I)(ii) and the relation of the CSA, there does not exist a 2^m -BFF design of resolution $R^*({1}|\Omega_3)$.
 - [2.2] When NSV₀ = 3 + p (p = 1, 2), we have $N \ge p + m + {m \choose 2} + {m \choose 3} \ge v_3(m)$ for $m \ge 6$ and p = 1, 2. Thus there does not exist a 2^m-BFF design of resolution R^{*}({1}|\Omega_3) with $N < v_3(m)$.
 - [3] When NSV₁ = 4, we have $N \ge 2m + {m \choose 2} + {m \choose 3} > v_3(m)$ for $m \ge 6$. Hence there does not exist a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_3)$ with $N < v_3(m)$.
- [c] When NSV₂ = 3, we have $N \ge 2\binom{m}{2} + \binom{m}{3} > v_3(m)$, and hence a 2^{*m*}-BFF design of resolution R^{*}({1}| Ω_3) with $N < v_3(m)$ does not exist.
- [C] When NSV₃ = $q \ge 2$, where $(m 3) 2 \ge q$, it holds that $\binom{m}{3} > 1 + m + \binom{m}{2}$ for $m \ge 5 + q \ge 7$. Thus we have $N \ge q\binom{m}{3} \ge 2\binom{m}{3} > v_3(m)$. Therefore there does not exist a 2^m -BFF design of resolution $\mathbb{R}^*(\{1\}|\Omega_3)$ with $N < v_3(m)$ for $m \ge 5 + q$.

Therefore the proof is complete.

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References

- Bailey, R. A. (2004). Association Schemes: Designed Experiments, Algebra and Combinatorics. Cambridge University Press, Cambridge.
- Bose, R. C., & Srivastava, J. N. (1964). Multidimensional partially balanced designs and their analysis, with applications to partially balanced factorial fractions. *Sankhyā: The Indian Journal of Statistics*, 26, 145-168.
- Chakravarti, I. M. (1956). Fractional replication in asymmetrical factorial designs and partially balanced arrays. *Sankhyā: The Indian Journal of Statistics*, *17*, 143-164.
- Ghosh, S., & Kuwada, M. (2001). Some estimable parametric functions for balanced fractional 3^m factorial designs. *Technical Report #01-6, Statistical Research Group*, Hiroshima University, Japan.
- Hyodo, Y. (1989). Structure of fractional factorial designs derived from two-symbol balanced arrays and their resolution. *Hiroshima Mathematical Journal*, *19*, 457-475.

- Hyodo, Y. (1992). Characteristic polynomials of information matrices of some balanced fractional 2^m factorial designs of resolution 2*l* + 1. *Journal of Statistical Planning and Inference*, 31, 245-252. http://dx.doi.org/10.1016/0378-3758(92)90033-O
- Hyodo, Y., & Yamamoto, S. (1988). Algebraic structure of information matrices of fractional factorial designs derived from simple two-symbol balanced arrays and its applications. In K. Matusita (Ed.), *Statistical Theory and Data Analysis II* (pp. 457-468). North-Holland, Amsterdam.
- Hyodo, Y., Yumiba, H., & Kuwada, M. (2015). Existence conditions for balanced fractional 2^m factorial designs of resolution $2\ell + 1$ derived from simple arrays. *Communications in Statistics-Theory and Methods*, 44, 2564-2570. http://dx.doi.org/10.1080/03610926.2013.788716
- Kuriki, S., & Yamamoto, S. (1984). Nonsimple 2-symbol balanced arrays of strength t and t + 2 constraints. *TRU Mathematics*, 20, 249-263.
- Kuwada, M. (1982). On some optimal 2^{*m*} factorial designs of resolution V. *Journal of Statistical Planning and Inference*, 7, 39-48. http://dx.doi.org/10.1016/0378-3758(82)90018-0
- Kuwada, M., Hyodo, Y., & Han, D. (2003). Characterization of balanced fractional 2^m factorial designs of resolution $R^*(\{1\}|3)$ and GA-optimal designs. *Journal of the Japan Statistical Society*, *33*, 181-201.
- Shirakura, T. (1975). On balanced arrays of 2 symbols, strength 2ℓ , *m* constraints and index set $\{\mu_0, \mu_1, \dots, \mu_{2\ell}\}$ with $\mu_{\ell} = 0$. *Journal of the Japan Statistical Society*, 5, 53-56.
- Shirakura, T. (1977). Contributions to balanced fractional 2^m factorial designs derived from balanced arrays of strength 2ℓ . *Hiroshima Mathematical Journal*, 7, 217-285.
- Shirakura, T. (1980). Necessary and sufficient condition for a balanced array of strength 2ℓ to be a balanced fractional 2^m factorial design of resolution 2ℓ . *The Australian Journal of Statistics*, 22, 69-74. http://dx.doi.org/10. 1111/j.1467-842X.1980.tb01155.x
- Shirakura, T., & Kuwada, M. (1975). Note on balanced fractional 2^m factorial designs of resolution $2\ell + 1$. Annals of the Institute of Statistical Mathematics, 27, 377-386. http://dx.doi.org/10.1007/BF02504657
- Shirakura, T., & Kuwada, M. (1976). Covariance matrices of the estimates for balanced fractional 2^m factorial designs of resolution $2\ell + 1$. *Journal of the Japan Statistical Society*, *6*, 27-31.
- Srivastava, J. N. (1970). Optimal balanced 2^m fractional factorial designs. In R. C. Bose, I. M. Chakravarti, P. C. Mahalanobis, C. R. Rao, & K. J. C. Smith (Eds.), *Essays in Probability and Statistics* (pp. 689-706). The University of North Carolina Press, Chapel Hill.
- Srivastava, J. N. (1972). Some general existence conditions for balanced arrays of strength t and 2 symbols. Journal of Combinatorial Theory, 13, 198-206.
- Srivastava, J. N., & Chopra, D. V. (1971). On the characteristic roots of the information matrix of 2^m balanced factorial designs of resolution V, with applications. *Annals of Mathematical Statistics*, 42, 722-734. http://dx.doi.org/10.1214 /aoms/1177693421
- Yamamoto, S., & Hyodo, Y. (1984). Extended concept of resolution and the design derived from balanced arrays. *TRU Mathematics*, 20, 341-349.
- Yamamoto, S., Shirakura, T., & Kuwada, M. (1975). Balanced arrays of strength 2ℓ and balanced fractional 2^m factorial designs. Annals of the Institute of Statistical Mathematics, 27, 143-157. http://dx.doi.org/10.1007/BF025 04632
- Yamamoto, S., Shirakura, T., & Kuwada, M. (1976). Characteristic polynomials of the information matrices of balanced fractional 2^m factorial designs of higher (2l + 1) resolution. In S. Ikeda, T. Hayakawa, H. Hudimoto, M. Okamoto, M. Siotani, & S. Yamamoto (Eds.), *Essays in Probability and Statistics* (pp. 73-94). Shinko Tsusho, Tokyo.

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