

Analysis of Liquidity-Study on Indian Mid-Cap Stocks

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Abstract

Liquidity is the pre-condition for a well-functioning and efficient market. Liquidity can be perceived, but difficult to measure due to its multi-dimensional characteristics. Studies have discussed various characteristics of liquidity and its influencing power on return and asset pricing. The article has considered Indian MidCap stocks and measured its liquidity using Amihud and trading volume as proxies. It has found Indian MidCap stocks have varying degree of liquidity. During the intraday, MidCap stocks have L-shaped liquidity pattern. The article observed that P-E ratio, P-B ratio, Dividend Yield and Index of Industrial Production are the significant determinants of liquidity. The article has estimated liquidity betas and carried out Granger non-causality test to articulate its relation with CAPM beta. The article has also found stability of liquidity beta across MidCap stocks. The liquidity betas of MidCap stocks have time-varying volatility. Relative Strength Index (RSI) and Change in Trading Volume are exogenous variables in explaining the time-varying volatility of beta. The study observed that MidCap stocks are claiming liquidity premium and liquidity premium is influencing the asset pricing along with WML, HML and EMR factors.

Keywords: liquidity proxies, liquidity pattern, liquidity premium, liquidity beta, carhart four factor model

1. Introduction

Liquidity of a stock is generally defined as the ability to trade large volumes with minimal price impact, cost and delay. While this definition itself is simple, a universally acceptable measure of liquidity continues to be elusive, resulting in the presence of diverse measures. Tightness, Immediacy, Depth, Breadth, and Resiliency are the primary characteristics of liquidity. The recent financial crisis highlighted the role of liquidity as a precondition for well-functioning and efficient markets. Findings about the properties of common determinants of liquidity will also help investors to decide on their liquidity exposures. With an improved knowledge of factors that influence liquidity, investor confidence will increase, leading to more efficient corporate resource allocation (Chordia et al., 2003). Literature highlights the role of liquidity in making various corporate and investment decisions (Gatchev et al., 2005). Lipson and Mortal (2009) examined the relation between equity market liquidity and capital structure and reported that firms with more liquid equity prefer equity financing. FII as a factor in determining trading activity and liquidity of stock markets is discussed in literature (Liu et al., 2009; Wang, 2007; Wei, 2010). A wide array of literature has studied the linkage between liquidity and stock returns in many countries (Faff et al., 2010; Narayan & Zheng, 2011). Investors are exposed to liquidity risk when the stock changes ownership making it a priced factor in making investment decisions (Jacoby et al., 2000; Acharya & Pederson, 2005; Nguyun et al., 2007).

The goal of this paper is to examine the liquidity of MidCap stocks listed on NSE using various liquidity measures; namely Amihud (2002) illiquidity ratio, Trading Volume, Spread and Depth. MidCap stocks are relatively more risky than large cap and less risky in comparison to small cap as investment options. Investors in emerging economies prefer MidCap stocks in their portfolio as with a liberalised industrial and services sectors, the chances of finding mid-level companies with potential to becoming large corporations are much better than the developed economies like the US and Europe.

With the above brief introduction, the study has initiated extensive review of literature with primary focus on the concept of liquidity, proxies and intraday patterns of liquidity, determinants of liquidity and its implications on asset pricing.

2. Literature Review

2.1 Liquidity Proxy and Pattern

According to Amihud and Mendelson (1986) a liquid markets tend to exhibit five characteristics: (i) tightness; (ii) immediacy; (iii) depth; (iv) breadth; and (v) resiliency. Tightness refers to low transaction costs, such as the difference between buy and sell prices, like the bid-ask spread in quote driven markets, as well as implicit costs. Immediacy represents the speed with which orders can be executed. Depth refers to the existence of abundant orders. Breadth means that orders are both numerous and large in volume with minimal impact on prices. Resiliency is the characteristic of the markets in which new order flow quickly to correct order imbalances. Market liquidity is impossible to capture with only one measure due to its multidimensional features.

An understanding of the behavior pattern of various liquidity proxies gives us an idea about the variations in the liquidity of stocks and the costs involved in trading in such stocks (Amihud and Mendelson, 1980; Acharya and Pedersen, 2005). Most of the existing literature gauging liquidity has focused on different dimensions of liquidity of individual financial assets. Baker (1996) concluded that there is no single unambiguous, theoretically correct or universally accepted definition of liquidity. Sarr and Lybek (2002) opine that there is no universally accepted measure to determine a market's degree of liquidity because of market specific factors and peculiarities. Bernstein (1987) examined different measures of stock liquidity and concluded that liquidity and efficiency are not compatible to each other. Amihud and Mendelson (1986) lay emphasis on the direct relationship between liquidity and cost of capital. High liquid markets are attractive to investors because of the easy exit from firm's ownership. Hui and Heubel (1984) hypothesizes that part of unsystematic risk represents liquidity of stock. Holden and Trzcinka (2009) conclude the Amihud measure does a better job than most other measures at capturing liquidity, and is robust to regime changes such as the change in minimum tick size to decimals. Illiquidity (ILLIQ) is much more precise as a measure of liquidity than other measures. Sioud et al. (2006) studied the dynamics of liquidity associated with the stocks listed on Tunisian stock market and concluded that shocks are absorbed more quickly for frequently traded stocks than for infrequently traded ones. Intraday pattern in stock markets has attracted much research attention. This importance is due to the existence of intraday regularities in stock market that contests the Efficient Markets Hypothesis. Köksal (2012) studies intraday patterns of various liquidity proxies on Istanbul Stock Exchange (ISE) using limit order book. This study reported spreads follow an L-shaped pattern whereas returns, number of trades and volume follow a U-shaped pattern. Tissaoui (2012) investigates the intraday pattern of trading activity, liquidity and return volatility in Tunisian Stock Exchange (TSE) and found U-shaped patterns. Krishnan and Mishra (2013) investigates intraday liquidity patterns on the National Stock Exchange (NSE) of India and reported that many liquidity proxies have U-shaped pattern.

2.2 Determinants of Liquidity

Jacoby and Zheng (2010) studied the empirical relationship between ownership dispersion and market liquidity. The study found that higher ownership dispersion improves market liquidity. Baber et al. (2012) studied the relationship between-Institutional Investors, Liquidity, and Liquidity Risk. They find that institutional ownership generally predicts larger stock liquidity, and so does concentrated ownership with mutual funds and hedge funds. Yaghoobnezhad et al. (2011) studied of relationship between Institutional Ownership and Stock Liquidity in Tehran Stock Exchange. The results of the study indicate a significant positive relationship between the percentage of institutional ownership and stock liquidity. Næs (2004) found a weaker negative relation between spreads and insider holdings in Norway stock exchange. Sharma (2005) studied ownership structure and stock liquidity on Indian stock market and found that the promoters' shareholding is not a statistically significant variable in explaining the determinants of liquidity in both Nifty stocks and Nifty junior stocks. Keim and Blume (2012) provide evidence that institutional participation in the U.S. stock market explains the cross-sectional variation in stock market illiquidity. Kim and Verrecchia (1994) study the relationship between earnings announcements, trading volume and liquidity. As per given model, earnings announcements increase the information asymmetry, which in turn leads to reduced liquidity in an imperfect market. Kumar et al. (2001) studied the impact of international listings like ADR and GDR on liquidity of Indian firm's underlying domestic shares. GDR listings are associated with enhanced liquidity while ADR listings (in most cases) are associated with reduced liquidity of the shares of domestic firm. Chordia, Roll and Subrahmanyam (2001) observed that liquidity and trading activity is influenced by market returns, its volatility, short-term and long-term interest rates. Macroeconomic news like GDP, unemployment rate also impact liquidity at the time of announcements. Gatchav et al. (2005) mentioned that the payout policy of the firm is related to the liquidity of its common stock. In illiquid markets investors will have greater demand for cash dividends from the stocks they hold. Larch and Peter (2011) sheds light on the actual impact of monetary policy on stock liquidity and thereby addresses its role

as a determinant of commonality in liquidity. Ding et al. (2013) empirically studied the relationship between Foreign Institutional Investors (FII) and stock market liquidity in China. The results reveal that with greater participation of foreign institutions stock market liquidity improves on both the Shanghai and Shenzhen exchanges. Goyenko & Ukhov (2009) document strong evidence for the U.S. monetary policy in influencing liquidity for the period 1962 to 2003. Chordia et al. (2005) report only modest predictive power of monetary policy for stock market liquidity. Söderberg (2008) studies the influence of 14 macroeconomic variables on the market liquidity of three Scandinavian stock exchanges between 1993 and 2005 and also provides mixed evidence.

2.3 Risk and Return and Liquidity

Amihud and Mendelson (1986) analyze the effect of bid ask spread or illiquidity on asset pricing i.e. how liquidity affects stock returns. Bali et al. (2013) revealed that stock market under-reacts to liquidity shocks on NYSE, AMEX and NASDAQ exchanges. Investor inattention and illiquidity both drive this under reaction. Bali et al. (2013) finds that immediate liquidity shocks have positive impact on contemporaneous stock returns. They examined double sorted portfolios using Fama-MacBeth regressions to confirm the highly significant relationship between liquidity shocks and future returns using large set of control variables example level of illiquidity, systematic liquidity risk, size, book to market, price momentum etc. Pastor and Stambaugh (2001) find evidence that market-wide liquidity is a key state variable for asset pricing on NYSE, AMEX and NASDAQ. Chang et al. (2009) analyzed the effect of liquidity on stock returns on Tokyo Stock Exchange (TSE) and found that the liquidity level has strong significant impact on stock return in different phases of business cycle. In case of Shanghai Stock Exchange (SHSE) and the Shenzhen stock exchange (SZSE), Narayan and Zheng (2011) found negative impact of liquidity on returns. Chordia et al. (2001) demonstrate the importance of trading activity related variables in the cross section of expected returns. Strong negative relationship is reported between both the level of liquidity, its volatility and expected returns. Chordia et al. (2001) argues that their finding is puzzling as risk averse investors require premium for holding volatile liquid stocks.

2.4 Liquidity and Assets Pricing

Acharya and Pedersen (2005) explained the economic significance of liquidity risk mentioned that liquidity risk explains about 1.1% of cross-sectional returns. Vu et al. (2014) examine the pricing of liquidity risk in the Australian market, using data from 1991-2010 and found strong evidence of co-movements of stock returns and market illiquidity and stock illiquidity and market returns. Hagstromer et al. (2013) investigates the relation between illiquidity and illiquidity risk and the size, value and momentum anomalies for US stocks. They found a very strong correlation between Fama-French size betas and illiquidity and a fairly strong correlation between Fama-French value betas and illiquidity risk betas. However, Carhart's momentum beta has high negative correlation with betas both for illiquidity level and risk. Eleswarapu and Reinganum (1993) empirically investigate the seasonal behavior of the liquidity premium in asset pricing and observed that liquidity premium is reliably positive only during the month of January. Martínez et al. (2005) empirically studied the explanatory power of systematic liquidity on asset pricing on Spanish stock market and concluded that market wide liquidity is a plausible factor to be included in asset pricing models. Hubers (2012) studies the relationship between asset prices and liquidity on London Stock Exchange (LSE) reported significance of liquidity betas.

3. Objectives of the Study

Looking at the literature review and above discussion, there is a requirement of formal approach to examine liquidity indicators and proxies, its determinants and implications of liquidity on asset pricing. The study focuses on following objectives:

- 1) Assessment of liquidity using both high and low frequency data at CNX Midcap stocks.
- 2) Estimation of liquidity beta and study its stability and causal relation with CAPM Beta.
- 3) Establish the relative importance of various firm specific factors and macroeconomic factors as determinants of liquidity.
- 4) Empirically examined the significance of liquidity premium as a factor in Fama-French asset pricing model.

4. Empirical Design

4.1 Proxy and Indicators of Liquidity

The study has used both low frequency day-end data and high frequency intra-day data to estimate following proxies for liquidity.

Low frequency proxies employed in the study are as follows:

- a) *Amihud (2002) illiquidity ratio*: This proxy provides a measure of price impact. The Amihud ratio is regarded as one of the best price impact measure constructed from low frequency data (Hasbrouck, 2006; Goyenko, Holden and Trzcinka, 2009). For individual stock, the illiquidity ratio ($ILLIQ_{id}$) is given by:

$$ILLIQ_{id} = \frac{|R_{id}|}{RVOL_{id}} \quad (1)$$

Where R_{id} is the return on stock i on day d and $RVOL_{id}$ is the Rupee volume of trading volume on the same day.

- b) *Trading Volume (Q_t)*: It is number of shares traded per time unit (e.g. day). Volume related measure captures the depth dimension of liquidity.

$$Q_t = \sum_{i=1}^{N_t} q_i \quad (2)$$

Where, N_t denotes the number of trades between $t-1$ and t , q_i is the number of shares of trade i .

High frequency proxies based on Chordia, Roll and Subrahmanyam (2000, 2001) are employed in the study. These are as follows:

- c) *Quoted Spread*

$$QSPR = |P_A - P_B| \quad (3)$$

Where P_A is the lowest ask price and P_B is the highest bid price.

- d) *Depth*

$$DEP = \frac{(Q_A + Q_B)}{2} \quad (4)$$

The study aims to discern and trace intraday liquidity movements in Indian stock market by studying high frequency data which is sampled at an interval of 30 minutes. This will help in understanding market microstructure in order to explain the existence of observed patterns.

4.2 Liquidity Beta

Following Watanabe and Watanabe (2008), the study will find out sensitivity of return to liquidity or 'liquidity β ' using the following regression equation:

$$R_i = \alpha_i + \beta_i Liq_i + \epsilon_i \quad (5)$$

Where R_i is the return, β_i is the liquidity beta and Liq_i is the liquidity of the stock i .

This equation is run for each stock on monthly basis using daily data and to obtain each stock's monthly β . For each industry, the stability of β will be tested through Augmented Dickey Fuller (ADF) test.

To verify the existence of a unit root in an AR (p) process, the test is $H_0: \beta = 1$ vs. $H_a: \beta < 1$ using the regression.

$$x_t = c_t + \beta x_{t-1} + \sum_{i=1}^{p-1} \phi_i \Delta x_{t-i} + e_t \quad (6)$$

Where c_t is a deterministic function of time index t and $\Delta x_j = x_j - x_{j-1}$ is the differentiated series of x_t . In practice, c_t can be zero or a constant or $c_t = \omega_0 + \omega_1 t$. The t ratio of $\hat{\beta} - 1$,

$$ADF = \frac{\hat{\beta} - 1}{std(\hat{\beta})} \quad (7)$$

Where denotes the least-squares estimate of β , is the well known Augmented Dickey-Fuller (ADF) unit root test.

The volatility clustering of 'liquidity β ' will be empirically examined using following ARCH model.

$$a_t = \sigma_t \epsilon_t \quad (8)$$

$$\sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2 + \dots + \alpha_m a_{t-m}^2 \quad (9)$$

Where $\{\epsilon_t\}$ is a sequence of independent and identically distributed (iid) random variables with mean zero and variance 1, $\alpha_0 > 0$, and $\alpha_i \geq 0$ for $i > 0$.

4.3 Determinants of Liquidity

As per the literature review, following macro-economic and firm specific factors would be used as explanatory variables.

- a) Index of Industrial Production (IIP)
- b) GSEC- 10 years yield (YLD)
- c) Net FII Inflow (NFII)

Firm Specific Factors:

- a) PE ratio (PE)
- b) PB ratio (PB)
- c) Dividend yield (DIY)

A panel data regression will be estimated using Amihud illiquidity measure as dependent variable and above mentioned factors as independent variables.

$$L_{it} = \alpha_{i0} + C_{i1}IIP_{t-k} + C_{i2}YLD_{t-k} + C_{i3}NFII_{t-k} + C_{i4}PE_{t-k} + C_{i5}PB_{t-k} + C_{i6}DIY_{t-k} + \epsilon_{it} \quad (10)$$

Where $t-k$ represents the lag length and ϵ_{it} represents the error term in regression.

4.4 Risk, Return and Liquidity Premium

Granger non-causality test is carried out on the time series data structure of proxy of liquidity against its returns. This is done in order to understand the causal relationship if any, between liquidity and return. Similarly, Granger non-causality test also carried out between liquidity beta and CAPM beta.

Stocks are sorted in decreasing order of liquidity and quantile portfolios are formed with one month formation and one month holding period. With two years data set of 30 companies, study has 36 portfolios each year. The return differential between the highest and lowest liquidity portfolio represents “Liquidity Hedging Portfolio (LHP)” return. This is as per the study of Fu et al. (2012). The study extended Fama-French four factor model by including LHP and empirically verify the same through the time series regressions. Identified sample stocks are sorted on the basis of market capitalization, book value, idiosyncratic volatility and returns. Sorting on the basis of market capitalization provides SMB factors. Similarly sorting on the basis of book value provides the HML factor. Momentum factor is estimated by sorting the portfolio on the basis of return. The return differential between the highest and lowest liquidity portfolio represents “Liquidity Hedging Portfolio (LHP)” return and it is the fifth factor. In addition to the four factors, Carhart (1997) model is augmented by a liquidity factor.

$$RP_{jt} - RF_t = a_j + b_j[RM_j - RF_t] + s_jSMB_t + H_hHML_t + m_jMOM_t + i_jLIQ_t + e_{jt} \quad (11)$$

Where, RP is the return for portfolio j , RF is the monthly risk-free rate; RM is the value-weighted market monthly return; SMB , HML , MOM and LIQ are the factor-mimicking portfolios for size, book-to-market, momentum and liquidity.

4.5 Data and Sample Selection

The study has selected 30 companies from the CNX–MidCap of National Stock Exchange of India. Mid-Cap stocks generally have liquidity problems in India. Stocks selection would be as per following criteria:

- a) Stocks which are highly priced and have high trading volume turnover should be excluded from the data set.
- b) Stock with very low price (pity stocks) and low trading volume should be excluded.
- c) Dataset should represent stocks from all sectors of the economy.
- d) Stocks should have 120 days of trading in a year.
- e) Stocks with missing values of dividend yield, book to market ratio, last periods annual returns, profit margin, leverage ratio, and earnings to price ratio are excluded.

The data set have representation of stocks from banking and finance, consumer goods, infrastructure, manufacturing, media, entertainment, healthcare, pharmaceutical, automobile and chemicals.

5. Results & Discussion

5.1 Liquidity Proxies, Patterns & Stability

Sector specific liquidity has been analysed using Trading volume and Amihud proxy. The estimation indicates

Pharma & Chemicals have the highest liquidity over the period and Consumer goods sector is the least liquid sector. This trend has been observed both in case of Trading volume and Amihud proxy (Figure 1 & 2).

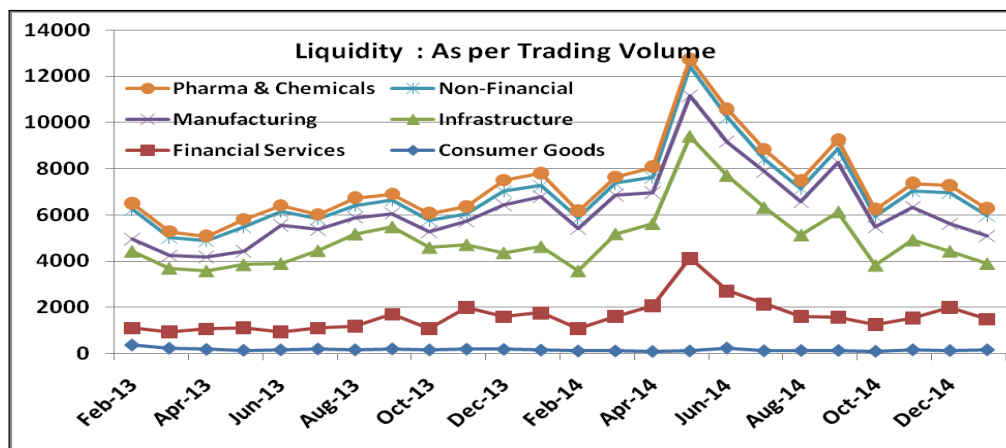


Figure 1. Liquidity as per trading volume

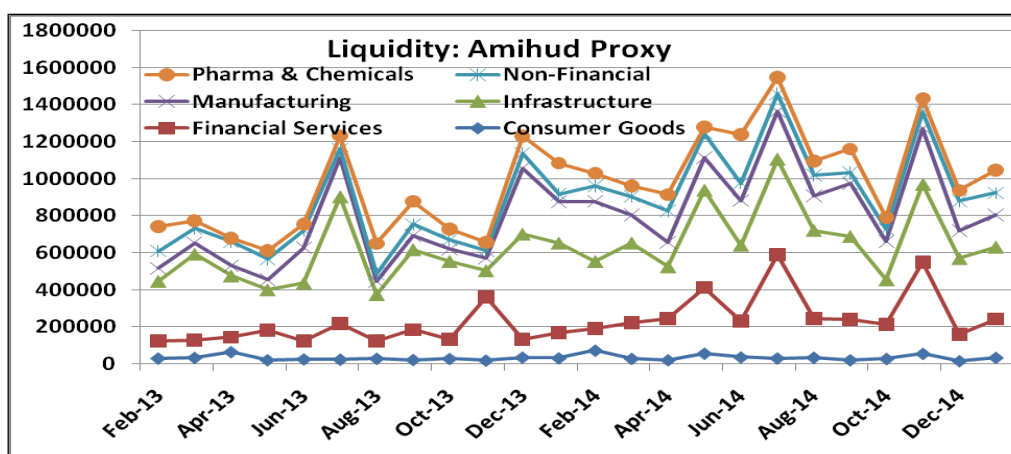


Figure 2. Liquidity as per amihud proxy

As per high frequency data (intraday Bid-Ask spread), the 30 stocks of MidCap indicates “L-shaped” pattern which has been found by other studies in different countries {Guo and Tian (2005) on Shanghai Stock Exchange; Köksal (2012) on Istanbul Stock Exchange}. The depth, as an indicator of liquidity, indicates varying degree of liquidity during the day.

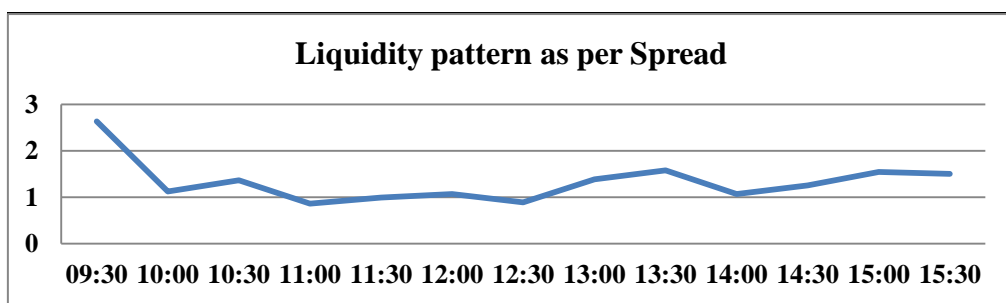


Figure 3. Liquidity pattern as per spread

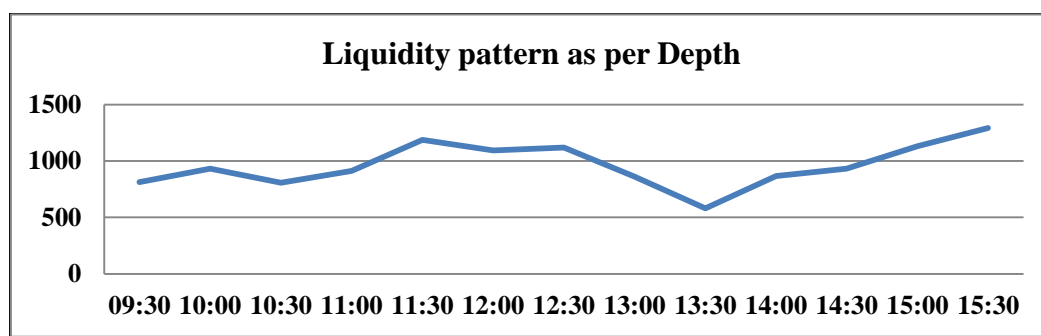


Figure 4. Liquidity pattern as per depth

Table 1. Mean Intraday spread of various sectors at an interval of 30 minutes

Time	Consumer Goods	Financial Services	Infrastructure	Manufacturing	Non Financial	Pharma & Chemicals
09:30	4.008	2.653	1.350	7.404	3.644	2.420
10:00	2.875	3.408	0.682	1.904	1.902	2.162
10:30	2.227	3.137	0.628	2.604	3.145	1.587
11:00	2.393	3.462	0.547	1.469	2.576	1.672
11:30	2.262	2.247	0.702	1.358	3.699	0.978
12:00	1.606	1.372	0.708	1.043	4.722	2.534
12:30	0.956	1.181	1.361	1.981	3.618	2.776
13:00	3.089	2.199	1.047	4.297	4.616	0.910
13:30	3.014	2.057	0.734	8.619	2.019	2.788
14:00	1.747	1.778	0.488	3.922	1.424	1.479
14:30	1.810	2.632	0.718	3.156	3.092	1.517
15:00	1.839	2.407	1.430	5.005	2.984	1.476
15:30	3.171	4.369	1.032	5.039	1.554	5.866

Table 2. Mean intraday depth of various sectors at an interval of 30 minutes

Time	Consumer Goods	Financial Services	Infrastructure	Manufacturing	Non Financial	Pharma & Chemicals
09:30	102	657	794	374	113	144
10:00	75	622	864	350	109	138
10:30	91	719	886	376	109	118
11:00	84	546	889	399	103	102
11:30	64	645	1116	488	169	110
12:00	73	619	1236	350	128	123
12:30	81	575	1114	371	113	123
13:00	81	523	1095	379	105	222
13:30	80	579	998	421	92	234
14:00	77	624	1114	296	100	141
14:30	173	655	1095	392	113	142
15:00	93	660	1175	424	151	113
15:30	112	697	989	1012	137	99

The study assumes that each sector of the economy should have stable liquidity structure. In order to verify this, sectoral stability of liquidity is examined using ADF test. The assumption is based on the fact that there will be continuous flow of information in different sectors of the economy. Also, this may not create unstable pattern of liquidity. The study has found unit root in case of Financial Services sector only, which indicates all other five sectors have stability in liquidity.

Table 3. Sectoral liquidity stability

Sector	ADF Test (P-Value)	Null Hypothesis: Has Unit a root (At 5% level of significance)
Consumer Goods	-6.3815 (0.000)	Rejected
Financial Service	-2.5591(0.1045)	Accepted
Infrastructure	-4.0754 (0.0016)	Rejected
Manufacturing	-2.8993(0.0485)	Rejected
Non-Financial	-2.60398(0.0951)	Rejected
Pharma & Chemicals	-4.06354(0.0016)	Rejected

The study has assumed that the volatility of liquidity will be time-varying as information passes through different time period. The time-varying liquidity volatility, for each sector, is examined through the ARCH effect with the null hypothesis of “no ARCH effect” is present. The Table No.2 indicates the presence of time-varying liquidity volatility in cases of Infrastructure and Non-Financial Sectors.

Table 4. ARCH effect of liquidity

Sector	LM Test (Probability)	Null Hypothesis: No ARCH effect is present
Consumer Goods	LM = 4.01275 $P(\chi^2(5) > 4.01275): 0.547582$	Accepted
Financial Service	LM = 2.89935 $P(\chi^2(5) > 2.89935): 0.7155$	Accepted
Infrastructure	LM = 18.3004 $P(\chi^2(5) > 18.3004): 0.00259$	Rejected
Manufacturing	LM = 6.08871 $P(\chi^2(5) > 6.08871): 0.29768$	Accepted
Non-Financial	LM = 20.5221 $P(\chi^2(5) > 20.5221): 0.00099$	Rejected
Pharma & Chemicals	LM = 0.288465 $P(\chi^2(5) > 0.288465): 0.99785$	Accepted

5.2 Liquidity Beta: Pattern, Stability & Granger Non-Causality

The liquidity beta is estimated for each of the 30 stocks on a monthly basis over the period of the study. The sectoral beta for each sector is computed using trading volume as weight. Figure 5, which depicts the pattern of liquidity beta for each sector indicates high variability of liquidity beta for Pharma & Chemicals and Non-Financial sectors.

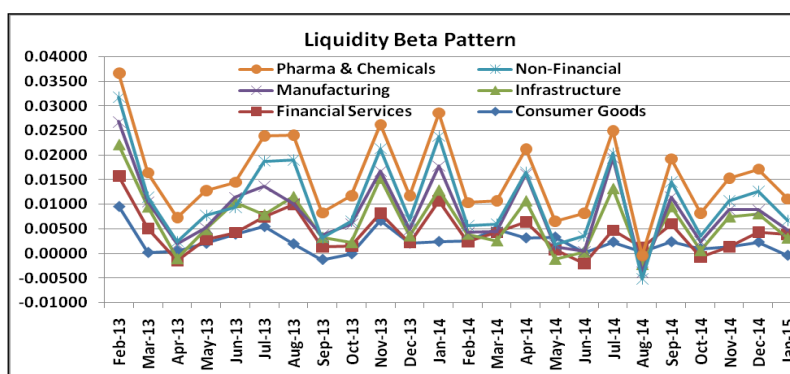


Figure 5. Liquidity beta pattern

The study assumes that each sector of the economy has stable liquidity beta structure. In order to examine the same, liquidity beta put through ADF test. All six sectors of the MidCap rejected the presence of unit root in liquidity beta, thereby indicating the stability of liquidity beta over the sample period (Table 5).

Table 5. Liquidity beta stability

Sector	ADF Test (Probability)	Null Hypothesis: Has Unit a root (At 5% level of significance)
Consumer Goods	-10.16517(0.000)	Rejected
Financial Services	-10.09439(0.000)	Rejected
Infrastructure	-9.097326(0.000)	Rejected
Manufacturing	-11.95518(0.000)	Rejected
Non-Financial	-10.83684(0.000)	Rejected
Pharma & Chemicals	-4.562242(0.0003)	Rejected

The CAPM beta is influenced by variability in Index returns, which in turn influences the trading volume and liquidity of different sectors. On the basis of the above fact the study assumes that the CAPM beta and Liquidity beta may have causal relationship. The Granger non-causality test indicates that neither CAPM beta nor Liquidity beta causes each other in cases of Financial Services, Manufacturing and Non-Financial sectors. Similarly, Liquidity beta causes CAPM beta in cases of Consumer goods, Infrastructure and Pharma & Chemicals sectors. However, in case of Pharma & Chemicals sector, the study doesn't find causality (Table No. 6).

Table 6. Granger non-causality (CAPM beta and liquidity beta)

Sector	CAPM Beta \nrightarrow Liquidity Beta	Liquidity Beta \nrightarrow CAPM Beta
Consumer Goods	F-Statistics: 0.7091(0.5874) Accepted	F-Statistics: 2.7897(0.030) Rejected
Financial Services	F-Statistics: 0.2760(0.8929) Accepted	F-Statistics: 1.985(0.102) Accepted
Infrastructure	F-Statistics: 1.03185(0.3943) Accepted	F-Statistics: 4.7478(0.0014) Rejected
Manufacturing	F-Statistics: 0.75197(0.5588) Accepted	F-Statistics: 0.6866(0.6028) Accepted
Non-Financial	F-Statistics: 0.41177(0.7998) Accepted	F-Statistics: 0.88911(0.4732) Accepted
Pharma & Chemicals	F-Statistics: 2.4807(0.0482) Rejected	F-Statistics: 2.45157(0.050) Rejected

The estimated betas of the 30 MidCap stocks, over the period of 24 months, are put through a GARCH test to understand their time-varying volatility. In the variance equation of GARCH, the study has included Relative Strength Index and Change in Trading Volume as explanatory variables to articulate the time-varying volatility of beta. The estimated GARCH equation (Table 7) indicates the presence of GARCH effect and the significance of both the explanatory variables. The robustness of the GARCH model is tested through residual normality test and ARCH effect.

Table 7. GARCH equation: Dependent: liquidity beta

Variable	Coefficient	Std. Error	Z-statistic	Prob.
Constant	0.002818	0.000269	10.49243	0.0000
Variance Equation				
Variable	Coefficient	Std. Error	Z-statistic	Prob.
Constant	2.33E-05	2.24E-06	10.39098	0.0000
RESID(-1)^2	0.149997	0.043523	3.446413	0.0006
GARCH(-1)	0.599995	0.065370	9.178426	0.0000
Relative Strength Index	-2.27E-07	2.02E-08	-11.21485	0.0000
Change in Trading Volume	-2.56E-06	1.21E-06	-2.116835	0.0343
Jaque-Bera: 570.0738(0.000)				
H0: No ARCH Effect: LM= 7.31541, P($\chi^2(5) > 7.31541$): 0.198219				

5.3 Liquidity Determinants

As per various literature (Ding et al., 2013; Kumar et al., 2001; Söderberg, J., 2008; Næs & Ødegaard, 2007), the study has considered P-E ratio, P-B ratio, Dividend Yield, Net FII, IIP and G-Sec 10 year yield as determinants of liquidity. The Amihud liquidity proxy of 30 MidCap stocks over the sample period are regressed, in panel data framework, against the above mentioned independent variables. As per the diagnostic tests (Table 8), panel model is adequate against pooled OLS and fixed effects are consistent as against random effects.

Table 8. Panel diagnostics test

Null Hypothesis	Test Statistic	Accepted/Rejected
Pooled OLS is adequate in favor of Fixed Effects Model	$F(29,684) = 20.888$ with p-value 2.358×10^{-75}	Rejected
Breusch Pagan Test:	LM=1318.73	Rejected
Pooled OLS model is adequate in favor of Random Effects Model	p-value [chi-square(1)] > 1318.73 = 9.605×10^{-289}	
Hausman Test:	$H=17.836$ with p-value = prob [chi-square(6) > 17.836] = 0.00665	Rejected
Random effects is adequate in favor of Fixed effects Model		

The fixed effects regression (Table 9) indicates P-E ratio, P-B ratio, Dividend Yield and IIP are significant determinants which influence the liquidity in the MidCap sector.

Table 9. Fixed-effects: Dependent variable-liquidity

	Coefficient	Std. Error	t-ratio	p-value
Constant	0.24825	0.897766	0.2765	0.78223
Price-Earning	-0.00205773	0.0010456	-1.9680	0.04947
Price –Book	0.125594	0.0251877	4.9863	<0.00001
Dividend Yield	-0.253616	0.0625278	-4.0561	0.00006
Net FII	-3.31421e-06	5.88994e-06	-0.5627	0.57383
IIP	1.88776	0.868354	2.1740	0.03005
GSec-10 Yr Yield	0.134618	0.107253	1.2551	0.20985
Adjusted R-squared: 0.48225				
F(35, 684): 20.134 (1.08×10^{-82})				
Durbin-Watson: 1.7845				
Null hypothesis:				
The groups have a common intercept				
F(29, 684) : 20.8879				
p-value: (F(29, 684) > 20.8879) : 2.25×10^{-75}				

5.4 Liquidity Premium & Asset Pricing

Stocks have been sorted in decreasing order of liquidity. The study constructed 72 portfolios with one month formation and one month holding period. The return differential between the highest and lowest liquidity portfolio represents “Liquidity Hedging Portfolio (LHP)” return. The Figure no.6 indicates the co-movement of LHP premium and portfolio return.

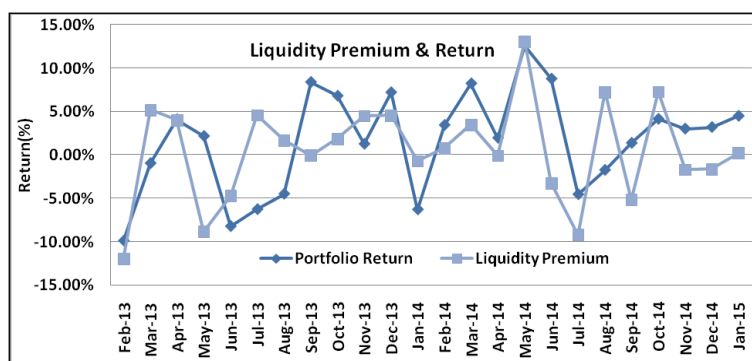


Figure 6. Liquidity premium and return

The study has also estimated premium for SMB (by arranging the stocks as per size), HML (by arranging the stocks as per Book value) and WML (by arranging the stocks as per momentum).

Table 10. Dependent variable: Excess portfolio returns

	Coefficient	Std. Error	t-ratio	p-value
Constant	-0.0252394	0.0148784	-1.6964	0.10704
LHP	-0.224064	0.119337	-1.8776	0.07674
SMB	0.180147	0.161769	1.1136	0.28010
HML	-0.392873	0.105789	-3.7138	0.00159
WML	0.288561	0.159168	1.8129	0.08655
EMR	0.843441	0.17741	4.7542	0.00016
Adjusted R-squared: 0.7228				
P-value(5,18): 22.624 (3.50e-07)		Durbin-Watson: 2.1283		

Following Carhart four factor model (1997), the study has estimated the asset pricing with LHP as the fifth factor. The estimated results (Table 10) indicate the significance of LHP, WML, EMR and HML (at 10% significance level) in asset pricing.

The studies of Chikore *et al.* (2014) and Sum (2014) empirically examined the causal relation between stock returns and liquidity. This study also tests the same in different sectors of the Indian economy (Table 11). The results suggest the acceptance of Granger non-causality between returns and liquidity across sectors except Financial Services where return Granger causes liquidity.

Table 11. Granger non-causality (return and liquidity)

Sector	Return \nrightarrow Liquidity	Liquidity \nrightarrow Return
Consumer Goods	F-Statistics: 0.85203 (0.4954) Accepted	F-Statistics: 1.40987 (0.2356) Accepted
Financial Services	F-Statistics: 3.16707 (0.0167) Rejected	F-Statistics: 2.42850 (0.0522) Accepted
Infrastructure	F-Statistics: 1.32060 (0.2670) Accepted	F-Statistics: 0.86689 (0.4864) Accepted
Manufacturing	F-Statistics: 0.15815 (0.9589) Accepted	F-Statistics: 1.39083 (0.2420) Accepted
Non-Financial	F-Statistics: 2.17787 (0.0763) Accepted	F-Statistics: 0.87763 (0.4800) Accepted
Pharma & Chemicals	F-Statistics: 0.99925 (0.4114) Accepted	F-Statistics: 0.69180 (0.5992) Accepted

6. Conclusion

Against the notion, the study indicates Indian MidCap stocks have varying degree of liquidity. Pharma & Chemicals sector has higher level of liquidity compared to consumer goods sector which is almost static during the studied period. The L-shaped for the intraday pattern indicates the declining trend of liquidity for the MidCap stocks with the business closing hours. The study has observed instability in the liquidity pattern of Financial Services Sector. Other five sectors of the MidCap have stable liquidity pattern. The study also observed that P-E ratio, P-B ratio, Dividend Yield and Index of Industrial Production are the major determinants of liquidity of MidCap stocks.

The study introduced the concept of liquidity beta and found that MidCap stocks returns are sensitive to liquidity variation. The study has also found stability of liquidity beta across the six sectors of the MidCap over the sample period. The study observed that CAPM and liquidity betas are independent in cases of Financial Services, Manufacturing and Non-Financial Sectors. The liquidity betas of MidCap stocks have time-varying volatility and Relative Strength Index (RSI) and Change in Trading Volume, as exogenous variables, in describing the time-varying volatility of beta. The study observed that MidCap stocks are claiming liquidity premium and also influencing the asset pricing along with WML, HML and EMR factors.

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