

The Monthly Effects in Chinese Gold Market

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Abstract

In this paper, we examine the calendar effects in Chinese daily gold returns from 30 December 2002 to 05 November 2011. We find that the average gold returns appear to be higher in February, September and November than in other months. Furthermore, using the GED-GARCH (1,1) approach, we find significant gold return anomalies in February, April, August, September, November and December. Most of these months occur before the public 7-day-holidays, which are called "Golden Weeks" in China. The expected strong demand before the holidays has pushed up the investment returns of gold. These findings provide convincing evidence of monthly effects in Chinese gold market. They might be of considerable importance for gold investors and traders.

Keywords: seasonal anomalies, calendar effects, gold price, Golden Weeks

1. Introduction

The monthly effect is one of the most frequently documented seasonal anomalies in capital markets. The returns in a specific month, e.g. January, are much higher than in other months. However, almost all previous studies have focused on the equity market. (Bouman & Jacobsen 2002; Fountas & Segredakis 2002; Lucey & Zhao 2008) It appears that very little attention has been paid on the calendar effect in the commodity market. Since gold plays an important role in hedging against the economic uncertainty in the equity market, it is worthy to investigate whether the monthly effect exists in the gold market. The deregulation of Chinese domestic gold market started from 2001. With the establishment of the Shanghai Gold Exchange (SGE) in 2002, the gold demand has been growing substantially. According to the "China Gold Report" (2010) by World Gold Council (WGC), China has ranked second in gold consumption, which amounts up to a combined total of 428 tonnes, and fourth in annual gold retail investment around the world in 2009. Furthermore, the Chinese gold consumption level is expected to double in the next decade. Apart from the gold's symbolic character of wealth in Chinese conventional culture, the rising resident income and domestic inflation expectation also pushed up the gold price. For instance, the closing spot prices of AU9995 has soared from 83.52 ¥ per gram in 30 December 2002 up to 338.16 ¥ per gram in 05 November 2012. Figure 1 illustrates the movements of daily closing prices (AU9995 and AU9999) from 2002/10/30 to 2012/11/5 in Chinese Gold market. We find in general the Gold prices keep rising. Due to the global financial crisis, the gold prices decreased in 2008. The average monthly prices of gold (AU9999) are shown in Figure 2, from which we find higher average prices occurred in May, August, September and October. In spite of some strong prices in these months, there has been no previous work that involves the calendar effects in Chinese gold market. This paper aims to contribute to the debate on calendar effects by examining the month-related seasonality of gold spot prices in Chinese market.

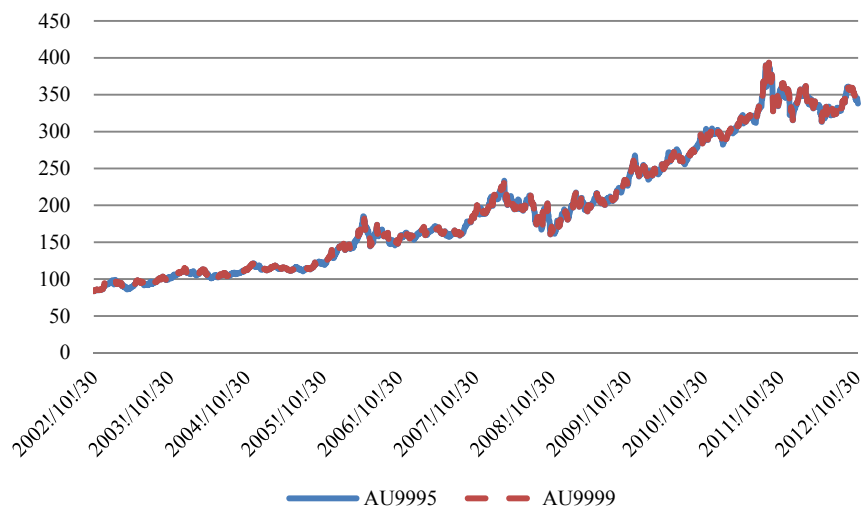


Figure 1. The movements of daily closing price in Chinese gold market

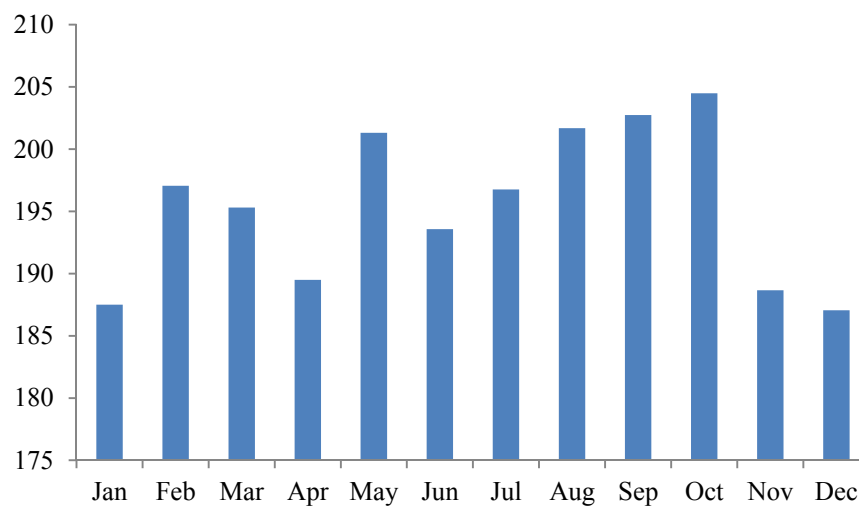


Figure 2. The average monthly price of gold (AU9999)

2. Literature Review

A large body of research has been devoted to the analysis of calendar anomalies and turn-of-month effects in stock returns. Odgen (1990) finds a surge in US stock returns around the turn of each calendar month and especially at the turn of the year. This result provides an explanation for the monthly effects and January effects in stock market. Using the ARCH model, Tong (1992) finds the January effect in the Taiwanese and South Korean stock market and discusses the Tax-Loss-Selling and liquidity constraint hypotheses. Mills and Coutts (1995) find the “January”, “Weekend”, “Half of the month” and “holiday” effects in the FT-SE 100, Mid 250 and 350 indices. Bouman and Jacobsen (2002) study the effectiveness of “Halloween Indicator” and suggest that the market returns between November and April are significantly higher in most countries, while during the period May-October are not significantly different from zero. Keef and Roush (2005) examine the Standard & Poor’s 500 stock index and find some unique features of the day-of-the-week effects in the pre-holiday returns. In terms of the emerging markets, Bepari and Mollik (2009) find an “April effect” in Bangladesh Dhaka Stock Exchange (DSE), while the “tax-loss-selling” hypothesis is not supported. Su et al. (2011) investigate the equal-weighted and value-weighted returns of the Chinese security markets. They find that the March effect, instead of January effect, in Chinese stock market. The delay of the turn-of-the-year effect can be attributed to the national

economical and cultural backgrounds. However, the calendar effect is also doubted by other researchers. Cheung and Coutts (1999) find no evidence of a January effect or any other monthly seasonality in the Hang Seng index. Marshall and Visaltanachoti (2010) suggest that implementing the Other January Effect (OJE) cannot bring excess risk-adjusted returns. Therefore the efficient market hypothesis cannot be rejected. Bohl and Salm (2010) also conclude that the OJE is not an international phenomenon. In terms of the commodity market, Girma and Paulson (1998) find the seasonality of petroleum future spreads and suggest that buy (sell) and hold trading strategies can be profitable. Lucey and Tully (2006) have used COMEX cash and futures data to investigate the seasonality of daily gold and silver contracts. They find a negative Monday effect in both gold and silver, across cash and futures markets. This finding also inspires us to investigate whether there exists the calendar effect in Chinese gold market. The structure of this paper is as follows. In section 2 we overview the existing literature on the concept of calendar anomalies. In section 3 we propose the methodology and the source of data. In Section 4, the results of the seasonality of gold prices are presented. Finally, we conclude the paper by summarizing the main findings.

3. Data and Methodology

We employ the daily closing spot prices of two primary gold products which are trading at the SGE: Au9995 and Au9999 (Note 1). Each price series runs from 30 October 2002 to 5 November 2012, providing 2425 end-of-day observations. All data is sourced from China Stock Market and Accounting Research Database (CSMAR) (Note 2). The prices are log-transformed to normalize the distributions. Following Mills and Coutts (1995), the continuously compound daily return of gold spot price (R_t) is defined as:

$$R_t = 100 \times \log(p_t/p_{t-1})$$

where p_t and p_{t-1} denote the gold spot prices at time t and $t-1$. Before estimating the calendar effect, we first test the stationary of both series of AU9995 and AU9999. The null hypothesis is unit root and the alternative hypothesis is level stationary. We use both ADF and Phillips-Perron (PP) including an intercept and a linear trend for the stationary test. The results are shown in Table 1. The unit root hypothesis cannot be rejected at the level for both series. But the first difference series are tested to be stationary.

Table 1. Unit root test results for both AU9995 and AU9999

	Level				First difference			
	AU9995		AU9999		AU9995		AU9999	
	ADF	PP	ADF	PP	ADF	PP	ADF	PP
None	1.84 (0.98)	1.79 (0.98)	1.85 (0.99)	1.79 (0.98)	-52.27*** (0.0001)	-52.28*** (0.0001)	-52.17*** (0.0001)	-52.16*** (0.0001)
Intercept	-0.31 (0.92)	-0.33 (0.92)	-0.30 (0.92)	-0.34 (0.92)	-52.35*** (0.0001)	-52.34*** (0.0001)	-52.25*** (0.0001)	-52.23*** (0.0001)
Trend & Intercept	-2.81 (0.20)	-2.87 (0.17)	-2.79 (0.20)	-2.85 (0.18)	-52.34*** (0.0000)	-52.34*** (0.0000)	-52.25*** (0.0000)	-52.23*** (0.0001)

Ordinary Least Squares (OLS) is a commonly used method to estimate the stock market seasonality. (Cheung & Coutts 1999; Keef & Roush 2005) However, the daily prices of financial assets normally exhibit the heteroskedastic and clustering features. (Schwert & Seguin 1990; Ball et al. 2006) The use of OLS leads to inconsistent parameter estimates because it assumes the constant variances (Note 3). To address this problem, Bollerslev (1986) and Taylor (1986) independently introduced the Generalized Autoregressive Conditional Heteroskedasticity (GARCH) model to allow for a time-varying conditional variance and richer dynamics. We use the following GARCH (1,1) model with dummy variables to estimate the month effect in Chinese gold spot market (Note 4).

$$R_t = \sum_{i=1}^{12} b_i D_{i,t} + \varepsilon_t \quad \text{with } \varepsilon_t | \Omega_{t-1} \sim N(0, \sigma_t^2) \quad (1)$$

$$\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2 \quad (2)$$

$$\varepsilon_t = z_t \sigma_t \quad \text{with } z_t \text{ is i.i.d and } z_t \sim N(0, 1) \quad (3)$$

Where R_t denotes the rate of gold returns. $D_{i,t}$ are the monthly dummies which indicate the month of the observation. In other words, $D_{1,t}$ takes a value of one if t is in January, otherwise it is zero. $D_{2,t}$ equals to one if t is in February and otherwise zero, and so on until December. We impose $\omega > 0, \alpha > 0$ and $\beta > 0$ in order to ensure the conditional variance (σ_t^2) is positive. If $\alpha + \beta < 1$, then the process ε_t is covariance stationary. z_t is a series of independently and identically distributed (i.i.d.) random variables with zero mean and unit variance. We use the Generalized Error Distribution (GED), proposed by Nelson (1991), to account for the leptokurtic behavior of our daily gold returns (Note 5). The probability density function of GED distribution is then defined as:

$$f(z_t; \nu) = \frac{\nu \exp[-0.5|z_t/\lambda|^\nu]}{\lambda 2^{(1+1/\nu)} \Gamma(1/\nu)} \quad (4)$$

where $\lambda = \left[\left(2^{-2/\nu} \Gamma(1/\nu) \right) / \Gamma(3/\nu) \right]^{1/2}$, $\Gamma(\cdot)$ is the Gamma function, ν is the thicknss-of-tail parameter (Note 6).

Before the estimation, we apply the BDS test to capture the non-linear structure in our sample. Under the null hypothesis that the series are i.i.d, the BDS statistic is asymptotically distributed as $N(0,1)$. The results of BDS tests on both return series are given in Table 2. We find that all the values of BDS statistics are high enough to reject the null hypothesis at the 1% level. Both time series of daily Gold returns are non-linearly dependent. Therefore, they cannot be explained by the linear model.

Table 2. BDS test statistics

Series	ε/σ	Value of ε	Value of m			
			2	3	4	5
AU9995	0.5	0.5847	7.73***	10.40***	13.18***	16.12***
	1.0	1.1695	9.21***	12.15***	14.70***	17.00***
	1.5	1.7542	9.94***	12.66***	14.74***	16.35***
	2.0	2.3389	9.92***	11.80***	13.33***	14.42***
	2.5	2.9236	9.93***	10.99***	12.02***	12.68***
AU9999	0.5	0.5789	7.94***	10.73***	13.57***	16.36***
	1.0	1.1577	9.06***	11.88***	14.33***	16.43***
	1.5	1.7366	9.58***	12.22***	14.30***	15.87***
	2.0	2.3155	9.58***	11.32***	12.97***	14.15***
	2.5	2.8944	9.33***	10.28***	11.50***	12.31***

Note: 2425 observations for each series, *** significant at the 1% level.

4. Results

Table 3 presents the results of monthly effect of daily Gold returns. We note that both return series of AU9995 and AU9999 have exhibited similar monthly effects. The average gold returns in February, September and November appear to be higher than in other months. This result is similar to the conclusion of “China Gold Report (2010)” which suggests that the strongest months for gold are January, September and November. Furthermore, the GED-GARCH(1,1) estimation results are positive and statistically significant in February, April, August, September, November and December, thus providing the evidence of monthly effect in Chinese gold market. We find that the gold returns in February, April, August and November are statistically higher than in any other month. These return anomalies are closely associated with the high demand of gold products before the holidays. Most of these months with high returns occur before the public 7-day-holidays, which are also “Golden Week” in China, i.e. the Lunar New Year in January/February (Note 7), Labor’s Day holiday in May and National Day holiday in October (Note 8). During these holidays, gold products, such as gold jewelries, decorations and coins, are decent gifts in Chinese traditions. Besides, the long public holidays are also popular and preferred time for young couple to hold their wedding ceremonies. All these factors contribute to the high demand of gold, especially the gold jewelries (Note 9). Furthermore, as pointed by “China Gold Report (2010)”, Chinese consumers typically restock gold during the winter months. As a consequence, the investment returns of gold are pushed up by the expected strong demand in February, April, September and December. These findings provide convincing evidence of calendar anomalies. The results have implied that all strong gold prices are associated with public holidays.

Table 3. Sample statistics and GED-GARCH (1,1) test results of daily gold returns

	AU9995			AU9999		
	Mean	Std.Dev	z-Statistic	Mean	Std.Dev	z-Statistic
January	0.0908	1.1300	0.1727	0.0961	1.1171	0.6805
February	0.1355	1.3556	2.4113**	0.1319	1.3409	2.1879**
March	-0.0517	1.0530	-0.2778	-0.0498	1.0662	-0.3412
April	0.0241	0.9931	2.5974***	0.0265	0.9716	2.6862***
May	0.0487	1.2909	0.9561	0.0481	1.2729	1.0870
June	-0.0494	1.1321	-0.1537	-0.0464	1.1146	0.0708
July	0.0356	0.9451	0.1757	0.0329	0.9281	0.0326
August	0.0885	1.1113	2.3710**	0.0870	1.0975	2.4318**
September	0.1224	1.4786	3.4419***	0.1250	1.4700	3.6972***
October	0.0023	1.4053	0.7265	0.0008	1.3965	0.1643
November	0.2099	0.9573	3.8672***	0.2049	0.9165	3.4876***
December	0.0542	1.1512	2.5680**	0.0434	1.1615	2.1165**

Note: 2424 observations for each series, ** significant at the 5% level, *** significant at the 1% level.

5. Conclusion

In this paper we examine the calendar effects in Chinese daily gold returns from 30 December 2002 to 05 November 2011. The daily returns of gold prices in February, April, September and December are proved to be statistically higher than in other months. The evidence of both AU9995 and AU9999 has supported this result. This phenomenon can be attributed to the “Golden Week” effects in China. The high demand before public 7-day-holidays has pushed up the gold spot prices. They might be of considerable importance for both gold investors and traders in Chinese gold market.

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Notes

Note 1. AU9995 is called “pure gold” and AU9999 is known as “thousand pure gold”, which has higher purity of gold than AU9995.

Note 2. CSMAR Database is designed and developed by GTA Information Technology, which is one of major providers of China data.

Note 3. We also estimated an OLS estimation with month dummy variables, and use the ARCH (1) test to check for the absence of heteroskedasticity. The test statistics (TR^2) for series of AU9995 and AU9999 are 79.93 and 77.49 respectively. Both values are high enough to reject the null hypothesis of no conditional heteroskedasticity.

Note 4. GARCH(1,1) model is appropriate to capture the volatility of a time series. (Hansen and Lunde 2005) It is also used in previous literature to investigate the seasonality and volatility structure of Gold and Silver future prices. (Lucey and Tully 2006; Batten and Lucey 2009)

Note 5. The assumption of a normal error distribution of a GARCH process is too restrictive for financial return series. In our sample, the skewness and kurtosis statistics of series AU9995 are -0.427556 and 10.20035 respectively. They are different from 0 and 3, which are the standard skewness and kurtosis values of Normal distribution. Similarly, the skewness and kurtosis statistics of series AU9999 are -0.409868 and 10.43266 respectively. Hence, we argue both series are leptokurtic and fat-tailed.

Note 6. For $\nu < 2$, the distribution has thicker tails than the standard normal distribution.

Note 7. In our sample, six out of ten Lunar New Year Holidays occurred in February. The others occurred in late January.

Note 8. However, from 2008 the Labor’s Day holiday has been reduced to three days.

Note 9. According to the WGC report (2009), jewellery contributed 78% of total gold consumption in China. It accounts for the vast majority of domestic gold demand.

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